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# Essays On CEO Compensation: New Evidence On The Managerial-Power

Vs. Optimal Contracting Debate

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#### Abstract

This dissertation analyzes the impact of an indicator of the likely future change in a firms' stock price – market analysts' expectations of share prices – on executive compensation. It examines how well the estimated effects of analysts expectations of share prices on compensation fit with two competing views of the determination of executive compensation: the arm's length bargaining model, which assumes that pay results from bargaining between the executives and a board of directors that seeks to maximize shareholders' interests, and the managerial power model, which assumes that the board and top executives seek to maximize managers' compensation within the constraint imposed by social costs and market penalties (Bebchuk et al. 2005). I focus on the pay of CEOs as the CEO is the most visible and important executive in a firm.

The first chapter documents the pattern of CEO pay from fiscal year 1996 to 2010. It uses the Execucomp data set of executive compensation reported on corporate financial statements, which is the most widely used data set for studies of executive pay and the Thomson-Reuters Insider Filing Data Feed (IFDF) on stock options and other derivatives paid to executives that firms report to the Security Exchange Commission. These two data sets complement one another. The Execucomp data set contains detailed information on cash compensation but only limited information on options and derivatives. The IFDF data contains the details of options but no information on salaries. I show that the number of options reported in the two datasets is highly correlated (except in 2006 when Execucomp changed the mode of reporting options), which justifies my combining them into a single data set. I find that CEO pay rose by \$2.7 million between 1996 and 2010, with most of the increase taking the form of greater stock-based compensation. I find a huge growth in stock options from 1996 to 2001 when the stock market had its dot.com boom. I also document that the percentage of options that were paid on a regular scheduled basis ("scheduled options") increased by almost 25% between 1997 and 2010 while the

practice of backdating options, which was common in the 1990s, disappeared.

The second chapter analyzes the Institutional Broker Estimate System (IBES) Detail History Price Target data file, a dataset that contains analysts' price targets for firms. This data has not previously been examined in terms of its relation to executive compensation. The price target data is an indicator of the likely future change in a firms' stock price. I show that the number of price target announcements issued by analysts is positively associated with company share price's volatility and that price targets are broadly predictive of changes in the value of the stocks. I also find that when analysts announce positive (negative) expectations of future stock price, share prices change in the same direction in the short run, which indicates that investors incorporate analysts' information in their assessment of firms. Given that the price target announcements are useful information about share prices, it is reasonable to expect that executives and boards will use them to help determine executive compensation. A positive price target should make stock options more valuable to an executive whereas a negative target makes salary compensation more valuable.

The third chapter analyzes the impact of price targets from the IBES Detail Price History Target database on executive compensation. I find that analysts' price targets alter the composition of executive pay between cash-based compensation and stock-based compensation. When analysts forecast a rise in the share price for a firm, the compensation package tilts toward stock-based compensation. When analysts forecast a fall in the share price, the compensation package tilts toward cash-based compensation. This pattern is more readily explicable by the managerial power model than the arm's length bargaining model of the determination of executive compensation. Consistent with the managerial power interpretation, the effect of analysts' price targets on the compensation package is stronger in companies that have weaker corporate governance.

The fourth chapter explores the impact of the introduction of the Sarbanes-Oxley Act (SOX) in 2002 and the Security and Exchange Commission's implementation of the Act in 2006 on the options granting process. I show that the introduction of SOX and its implementation educed and

eventually eliminated the practice of backdating options (giving options at time t but dating them at some earlier time when the share price was lower). But it is also associated with greater "spring loading" of option grants around analysts' price targets announcements (timing the granting options so that executives receive options after bad price announcements or before good price announcements). An increased proportion of options were granted around the time of analysts' price targets announcements from 1996 to 2010 while the number of options granted independently of the announcements fell. I argue that an ideal contracting board and a rent-seeking managerial power board will use the information in the future expected share price differently in setting executive pay. The shift from backdating to spring loading after enactment of SOX and the relation between options and the analysts price targets are more readily explicable by the managerial power model than the arm's length bargaining model.

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#### **General Introduction**

Since at least the 1980s inequality of income in the US has been high and rising (see figure 1). Part of the increase is associated with a huge jump in the ratio of top executive compensation to the compensation of ordinary workers. At this regard, Schlozman *et al.* (2005) found that Americans consider certain jobs, such as CEOs, overpaid. Lazonick (2010) argues that the ballooning compensation of executives is the main cause behind the increase in inequality in U.S. The way CEOs are paid has attracted the attention of the press as well as the one of the academic literature, becoming one of the hot topics in the last years.





Academic views about CEO and other high level executive compensation fit into two groups. On one side, some researchers believe that the inequality arising from top executive compensation package is efficient. Independent boards of directors set incentive schemes

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(bonuses and stock-based compensation such as stock options) for executives to align managements' interest to those of the shareholders, which ultimately maximizes shareholder's interests. On the other side, some analysts argue that managers use private information to structure compensation for their benefit at the expense of shareholders. Management appoints persons to the board of directors who are likely to weigh management heavily in setting compensation so that rather than solving principal-agent problems, incentive schemes are part of the agency problem itself.

Bebchuk et al., (2003) explore how the two models influence top executive compensation. They argue that the observed pay structure is most likely a compromise between market forces, which mitigates managerial rent-seeking, and managerial-power that favors top executive compensation. Murphy (2002) questions the ability of both the arm's length bargaining model and the managerial power model to fit the observed pattern of executive pay. He shows that both models can explain the rise in CEO's compensation under different assumptions but highlights inconsistencies between the evidence and the managerial power model, such as the coincidence of increased CEO pay and greater board independence; the attractive compensation packages awarded outsiders to join top management. He favors a "perceived-cost view" - that assumes that risk adverse executives cannot hedge the risk of options (which implies that the Black-Scholes formula overvalues them<sup>1</sup> and that company's underestimate the cost of granting options to executives, which yields conclusion similar to those offered by the managerial power model but with different policy implications (Murphy, 2002; Hall et al., 2003). Reviewing empirical studies, Frydman et al. (2010) conclude, "both managerial-power and competitive market forces are important determinants of CEO pay" but leave open the question of their relative importance.

<sup>&</sup>lt;sup>1</sup> The Black Scholes formula commonly used to evaluate the stock options value assumes that the holder can sell the options or hedge the risk. Executives, however, cannot diversify their portfolio fully invested in the company and thus their options are overvalued.

My dissertation contributes to the existing debate over which model better explains CEOs compensation and ultimately contributes to understanding the part of the rise in inequality in U.S. associated with increased executive pay. Specifically, my thesis examines how the two main models of the determination of executive pay – the arm's length bargaining model and the managerial power model- account for estimated relations between expected share prices and compensation. I analyze the impact of an indicator of the likely future change in a firms' stock price on executive compensation. The indicator is the expectations that market analysts give for the firms' share price in the next year as reported by the Institutional Broker Estimate System in its Detail Price History Target data base. The share price that analysts expect for a firm in the future has not, to my knowledge been previously analyzed in studies of executive compensation. I show that the way in which it affects the composition of executive compensation provides insights into the determinants of executive pay.

The thesis has four chapters.

Chapter one documents the changes in the ways in which firms compensated CEOs from fiscal year 1996 to 2010, as given in the widely used Execucomp data set that obtains information on executive compensation from corporate financial statements and as given in the less widely used Thomson-Reuters Insider Filing Data Feed (IFDF) on stock options and other derivatives paid to CEOs that firms report to the Security Exchange Commission.

Chapter two analyzes the Institutional Broker Estimate System (IBES) Detail History Price Target data file, a unique, comprehensive and richly detailed data set that portrays analysts' price targets, an indicator of the likely future change in a firms' stock price. Focusing on US firms only, I show that the number of price targets announcements issued by analysts is positively associated with company share price's volatility and that price targets are informative about future share prices and that positive (negative) announcements are followed by an increase (decrease) in share price shortly afterward, which implies that investors incorporate analysts' information and adjust share prices according to announcements. Chapter three analyzes the impact of price targets, as reported by the Institutional Broker Estimate System in its Detail Price History Target data base, on executive compensation. I find that analysts' price targets alter the composition of executive pay between cash-based compensation and stock-based compensation. When analysts forecast a rise in the share price for a firm, its compensation package tilts toward stock-based compensation. When analysts forecast a fall in the share price, the compensation package tilts toward cash-based compensation. This pattern is more readily explicable by the managerial power model than the arm's length bargaining model. Consistent with this interpretation, the trade-off is stronger in companies that have weaker corporate governance.

Chapter four explores the impact of the Sarbanes-Oxley Act in 2002 and the Security and Exchange Commission's implementation of the law in 2006 on the options granting process. I show that the new law essentially eliminated backdated options but led firms to substitute spring loading options by timing them strategically around analysts' price targets announcements (forward-looking strategy). I test which model (length-arm bargaining model vs. the managerial power model) fit most the results and I found that the evidence is more readily explained by a rent-seeking model.

#### Chapter One: New Evidence on CEOs' Compensation

Over the last decades, the magnitude and mode of compensation of CEOs in the US has changed greatly. In this chapter I document the changed way in which firms compensated CEOs from fiscal year 1996 to 2010. I combine data from the Execucomp data set<sup>2</sup> that obtains information on executive compensation from corporate financial statements with data from the Thomson-Reuters Insider Filing Data Feed (IFDF)<sup>3</sup> on stock options and other derivatives paid to CEOs that firms report to the Security Exchange Commission. Execucomp is the most widely used data set in the literature on executive compensation. It contains information on all components of executive pay but has only limited information on salary or other forms of complete information on the details of options but not information on salary or other forms of stock options. To see whether the two data sets are sufficiently comparable to be melded into a single data set for analysis, I compare the one variable that they both report - the number of options reported in the two datasets but they are sufficiently highly correlated to support my combining them for some analysis.

Before presenting the data I review briefly the debates over the determination of executive pay and, in particular, the role of stock-based pay as a mode of aligning the incentives of executives with the interests of shareholders that guides my analysis of the changes in compensation over time.

<sup>&</sup>lt;sup>2</sup> Execucomp collects top executive compensation detailed information directly from company's annual proxy statement (DEF 14A SEC).

<sup>&</sup>lt;sup>3</sup> The Insider Filing Data Feed (IFDF) reports U.S. insider activities filled on Form 3, 4, 5, or 144. Table II capture official grant date, the exercise date, the expiring date and the exercise price of any insider's derivative grants.

#### 1. Introduction

Proponents of the principal-agent model of pay have stressed the ways in which a board of directors can institute incentive schemes to reduce the agency costs from separation of ownership and control and improve shareholder value. Ideally, such compensation systems align executives' interests to those of shareholders (Jensen *et al.*, 1976; Fama *et al.*, 1983; Eisenhardt, 1989; Dalton *et al.*, 2007). Consistent with such arguments, firms increased stockbased compensation considerably in the last decades.

A board that seeks to motivate executives with different forms of pay should equate the marginal cost to the firm of additional option or stock or other form of compensation to the extra benefits from inducing executives to make decisions in the shareholders' interests. Adherents of the view that executive compensation has a substantial positive impact on economic performance of firms believe that performance depends critically on managerial skills, and that boards negotiate stock-based compensation to create incentives for a manager to raise the share price, which they take as the best measure of the long term value of a firm. More able or skilled managers are more likely to accept stock-based compensation than less skilled managers.

Critics of executive compensation argue that because executives influence membership on a board that boards make decisions about compensation that benefit executives at the expense of shareholders, and/or that executives find ways to game whatever incentive system the board sets to enrich themselves. Bebchuk *et al*, (2005) notes that the increase in equitybased compensation "has not been accompanied by a substitution effect, i.e. a reduction in nonequity compensation", which suggests that the stock-based incentive system may simply be a new mode for executives to enrich themselves. But without evidence on the true marginal value of an executive it is also possible that their value increased over time and that firms increased equity-based compensation in line with increased value. The US system of basing executive options on the share price of a firm rather than on the share price of the firm relative to that of competitive firms or to the stock market broadly is also suggestive of a system run for the benefit of executives. Under the US system, management can profit from fortuitous increases in the firm's share price -good luck- that have nothing to do with their effort (Bertrand *et al.*, 2001), by general inflation in share prices, or by using their inside information to time the receipt of options in ways that would most benefit them.

Summarizing the debate, Bebchuk *et al.* (2003, 2004) divide adherents and critics of executive compensation into two opposing camps. On the one side are adherents to "the arm's length bargaining model" that posits that boards of directors try to maximize shareholders' interests subject to the constraint set by executives' supply behavior -paying the amount and nature of compensation that is necessary for potential executives to accept a job and to spend their time and effort finding ways to increase the value of shares. The package is the result of arm's length transaction between the board seeking the best deal for shareholders and executives selling their skills. On the other side is the "managerial-power" model, which assumes that the board seeks to maximize managers' compensation within the constraints imposed by social costs and market penalties. Under the managerial power model the board of directors seeks the most favorable compensation for executives under the threat of market penalties and social costs that rise when it adopts excessively favorable arrangements for executives (Bebchuk *et al.*, 2005).

Bebchuk *et al.* (2005) find that both models help explain the increase in top executives' compensation. In their recent review of the executive compensation literature Frydman *et al.* (2010) reach a similar conclusion: "both managerial-power and competitive market forces are important determinants of CEO pay, but that neither approach is fully consistent with the available evidence."

The increases in executive pay, public concerns over that pay, and the existing of two competing schools of thought have made executive compensation a hot area among academic researchers. I typed the keywords "CEO compensation" in Pro-Quest Central, a database of periodical content featuring scholarly journal, trade publications, newspaper, magazines and other reports and found 315 scholarly articles in 1996. Ten years later the number of article reported more than doubled to 765. In 2010 the articles reached 971. Executive compensation has also attracted increased public interest.<sup>4</sup>

year	Pro-Quest- Total	Pro-Quest- Newspaper	Pro-Quest- Scholarly Journals	Pro-Quest- Magazine	Google Scholar Pages
1996	4891	2462	315	250	1250
1997	5652	2736	352	281	1390
1998	8022	3210	432	350	1770
1999	8165	2886	467	311	2280
2000	9574	3220	430	278	2760
2001	10285	3155	559	277	3370
2002	14671	5405	586	499	4150
2003	16643	6129	712	599	5180
2004	21375	5563	733	637	5630
2005	32837	5330	816	603	6820
2006	44292	6287	765	638	7610
2007	45426	6576	827	554	8600
2008	43044	6072	814	503	8970
2009	36544	6739	865	507	10500
2010	33538	6542	971	344	10400

Table 1.1: Articles on CEO compensation in Pro-Quest Data-Base.

The number of newspaper articles on CEO compensation in Pro-Quest Central increased from 4,891 in 1996 to 33,538 in 2010. The number of magazine articles in Pro-Quest Central relating to executive pay went up from 250 to 344 over the same period. Similarly, Google Scholar gives 1250 pages in which the words "CEO" and "compensation" appear in the title of articles contained for 1996, and 10,400 pages fourteen years later.

The increase in academic and popular media articles on CEOs' compensation reflects the concern of American people that CEOs may be overpaid and wider concerns about the distribution of income. Schlozman *et al.* (2005) find that Americans are critical on CEO pay level.

<sup>&</sup>lt;sup>4</sup> Kuhnen et al. (2012) explore the impact of public opinion on executive compensation.

The academic research and public concern motivate my documenting in this chapter the changes in CEO compensation during the 1996-2010 period. My analysis of the data shows that CEO pay rose during the dot-com bubble to reach a maximum in nominal term in 2000, after which it decreased until 2004 and then began increasing again. The components of compensation also changed in this period. There was a huge increase in the share of executive compensation that took the form of stock-based compensation. Stock options increased to become the largest single component of compensation in 2000/2001. The vast majority of options granted to CEO were at the money with the vesting period centered around one to three years. The practice of backdating options – in which firms give an option at time t but report that it was given in t-1, when the price of shares was much lower, proliferated in the 1990s, then fell after the 2002 Sarbanes-Oxley Act strengthened reporting requirements. Backdating disappeared when the SEC strengthened the reporting rules so that backdating became nearly impossible to carry out. The data also show a substantial drop in the use of options in the late 2000s, with compensation shifting to restricted stock grants and to non-equity incentives, possibly as an alternative to stock options, for aligning the interests of executives and shareholders (arm's length bargaining model) or as an alternative way to enrich executives (managerial power model).

I describe next the sources of data that I use to measure the changing level and composition of executive pay and note the strengths and weaknesses of the data.

To construct my dataset for analyzing executive compensation I use three different sources of data. First, I obtained data on the components of top executive compensation, like salary, bonuses, restricted stock and value of stock options granted from the Execucomp files of Compustat. This is the most widely used source of data on top executive compensation in academic studies. But because it does not provide much detail on the stock options that have become such an important part of executive compensation, I went to a second less widely used data set: the Thomson Financial Insider Filling (IFDF) database that provide detailed information regarding stock options granted to CEOs, such as grant date, exercise price, expiration date, number of derivatives and the number of shares granted to insiders from 1996 to 2011. Finally, to evaluate options and to measure whether they are backdated or otherwise granted in ways that benefit executives independent of the incentive effects, I obtained firms' daily stock price from Center for Research in Stock Price (CRPS).

#### 2. Execucomp data files

Standard & Poor's Executive Compensation data set (Execucomp) collects top executive compensation data from company's annual proxy statement (DEF 14A SEC) for the S&P 1500 companies and others. I use the data for fiscal year 1996 through fiscal year 2010. In addition, I have obtained data about the firms from Compustat.





Figure 1.1 shows the amount of nominal dollars in each component of compensation for CEOs in the Execucomp files. The figure stacks the data for each form of pay, and thus provides a picture of both the level and composition of CEO pay from fiscal year 1996 to 2010.

The top of each bar measures the average level of total compensation. The different elements show the average level of each component: salary, bonus, stock options, shares, long term incentive payout, deferred compensation, non-equity incentive plan and other forms of compensation. Total CEO compensation increased from a bit more than \$3 million in 1996 to \$5.7 in 2010. The increase was very large from 1996 to 2000 during the dot.com boom. Total compensation reached a maximum in 2000, and then declined through 2004 after which it started to rise again. However, nominal compensation in 2010 was still lower than in 2000.

The increase in executive compensation over the entire period took the form mainly of an increase in stock-based compensation. In 1996 stock-based compensation constituted 55% of total compensation (calculated as the sum of salary, bonus, options, stock, LTIP, deferred compensation, non-equity and others compensation). In 2000 stock-based compensation was nearly 74%. Most of the increase in stock-based compensation took the form of increased stock options so that by 2000 stock options constituted approximately 67% of total compensation. Thereafter, stock-based compensation shifted from options to restricted stock grants. In 2010, stock grants constituted about 35% of total compensation. Salaries, by contrast, increased more modestly – from \$550,000 in 1996 to \$822,000 in 2010. The salary share of total compensation fell from 18% of total compensation to 14% of total compensation.

Because the Execucomp sample of firms changes over time, it is possible that the averages in figure 1.1 are distorted, at least in terms of firms' experiences, by the changing composition of the sample. To see how salaries of an unchanged sample of firms changed over the period, I took firms which appeared in consecutive years of Execucomp and calculated the yearly change in total compensation year by year. For instance, if a firm in Execucomp reported total CEO compensation of \$992,722 in 1997 and total CEO compensation of \$1,176,401 in 1998, I computed the difference between the two years. I averaged the differences across all the firms that had data on compensation in both years. Table 1.2 below gives the average of the differences in thousands of dollars.

year	Change for fixed sample over previous year	Cumulative Change	
1996-1997	794		
1998	834	1628	
1999	799	2427	
2000	1794	4221	
2001	-566	3655	
2002	-1139	2516	
2003	-354	2162	
2004	648	2810	
2005	239	3049	
2006	127	3176	
2007	15	3191	
2008	-168	3023	
2009	-563	2460	
2010	950	3410	

 Table 1.2: Average year-to-year changes for same companies in total compensation in thousands of dollars.

The cumulative sum of all the changes in the line for 2010, \$3,410, estimates the change in average pay for the sample of firms that were unchanged in each year. This number exceeds a change in average executive compensation from 1996 to 2010 for the sample of firms that reported in 1996 and 2010 of 2.7 million. The higher estimated change from year to year presumably suggests that firms appearing in consecutive years are reasonably successful and thus more likely to have steadier increases in pay than the sample of firms whose identify changes from year to year. But the overall pattern of change mimics closely the changes shown in figure 1.1. Since the cumulative sum in table 1.2 peaks in 2000, this calculation confirms the finding that total compensation reached a maximum in 2000.

There is an important change in the Execucomp data in 2006, when FAS123R changed the reporting rules for company reports of stock options. Before 2006, Execucomp provides its own fair award value estimates of the monetary worth of stock options granted using a nonstandard Black-Scholes formula. After 2006 companies had to report the estimated fair value of the stock options granted in the proxy statement, and Execucomp decided to drop its method in favor of the company reports. The SEC allows companies to use different methods of evaluating options, including the Black-Scholes and the binomial options pricing model. The result is that the data from 1966 to 2005 value stock options granted on the basis of Execucomp's analysis, whereas the data from 2006 to 2010 value stock options granted on the basis of the evaluations by each company.

The reason Execucomp used a non-standard Black-Scholes formula to value stock options and the SEC allows firms to use their own methods of evaluation is that the Black-Scholes formula is designed for European Options, which can be exercised only at the expiry date, whereas almost all options traded in America are American Options, which can be exercised any time up to the expiration date (Bodie et al., 2001).<sup>5</sup> Since holders of the options have greater leeway over when to exercise them they generally tend to have higher value than the Black-Scholes formula. Indeed, empirical analysis shows that the Black-Scholes formula tends to undervalue options deeply in the money but overvalue call options deeply underwater (Bodie et al., 2001). Geske et al. (1984) give evidence that this result is due to the fact that Black-Scholes model does not take into account the opportunity to exercise American options earlier than European ones. Regarding the letter point, Compustat computes the Black-Scholes value of the options by using 70% of the stated life. For instance, if the time to maturity is 10 years, Compustat reduces it to 7 years. Whaley (1982) shows that more complicated models do better than Black-Scholes by taking into account early exercise. However, Bodie et al. (2001) stress that many empirical studies show that Black-Scholes is a fair approximation of traded options actual prices.

There is a further change in the SEC's compensation disclosure rules adopted in 2006 that affects the data. Before 2006, there was a distinction between short-term incentive pay and long-term incentive pay (LTIP). In 2006 the SEC introduced a new terminology to isolate

<sup>&</sup>lt;sup>5</sup> Also, most options granted to CEOs can be exercised only after a certain period of time (vesting period) as shown in table 1.4. They are referred as "Bermudian" Options since they are a hybrid between American Options and European Options (Rubinstein, 1995).

executive incentive compensation paid in cash: non-equity incentive. The new category encloses incentives awarded in cash that are earned if management meets certain performance criteria. Prior to this, firms reported all short-term incentive paid in cash as part of the bonus component of CEO compensation. Thus the new terminology shifted of cash compensation previously reported under the label "bonus" to the new label "non-equity incentive". Hence, thereafter 2006, "bonus" defines cash earned by officers who met criteria that were not disclosed, while it no longer reports LTIP.





Figure 1.2 illustrates the evolution of CEO compensation with the modes of compensation regrouped for consistency over time by putting bonuses, long-term incentive payout and non-equity incentive under a single label. This shows that incentives paid in cash increased from 23% to 27% from 1996 to 2010.

#### **3.** Thomson-Reuters Financial Insider Filling Database

As noted, Execucomp does not provide details on the growing proportion of pay that is

stock-performance based. It has only limited evidence on the share options that became so important in the late 1990s through the mid 2000s. To obtain detailed data on options granted to executives I turn to the Thomson-Reuters data set, and in particular to Insider Filing Data Feed (IFDF) Table II. From this table I obtain the official grant date, the exercise date, the expiry date and the exercise price of insider's grants, which allows me to infer if the options were scheduled (given at regular intervals from year to year) or unscheduled (given at irregular times); were backdated; and if the options were granted at the money, out of the money, or in the money. IFDF also provides the vesting period and the maturity period. In the data appendix 1.A I give detail of the methodology adopted to construct my dataset.

I focus on transaction involving CEOs for the fiscal year from 1996 to 2010. Table 1.3 summarizes statistics for CEO receipt of options from the IFDF Table II. The first three columns show the fiscal year of the transaction, the number of firms in the data, and the number of derivatives granted. The remaining columns show the characteristics of the options granted in terms of the percentages that fit into different categories. Column 4 shows that options are generally granted at the money but that the proportion granted at the money has increased from 77% of all options in 1996 to almost 96% in 2010.

Column 5 reports the percentage of derivatives that I infer were backdated. I give the full detail of how I estimated the number of backdated options in the appendix. The basic idea is that whenever firms report an excessively large number of options on days when the share price is especially low, they are likely to be backdating. If for instance firms can give options on three days, it is reasonable to expect 1/3<sup>rd</sup> to be given on each day. If all the options are given on the day with the lowest share price, that would be a sign of potential backdating. The column shows a drop in inferred backdating after 2002. The most likely reason for this is the enactment of the Sarbanes-Oxley Act of 2002, which imposed that companies notify within two business days a change of ownership of officers' options. There is another drop in the percentage backdated in 2006 and 2007. The likely reason for this is that in 2006, the SEC

strengthened the 2-day rules in an attempt to further reduce backdating (Bickley *et al.*, 2008). In my analysis using a window of 41 days, the chance that an option would be granted at the lowest value period within the option is 1/41 (2.44%). The average "% estimated backdating" from 2007 to 2010 is 1.96%, which is a bit below though not statistically different from 2.44%, so I conclude that backdating has essentially ended.

Year	Firms	Transaction	% At-the- Money	% Estimated Backdated	% Estimated Scheduled	% Estimated Unscheduled	% Unclassified
1996	728	1455	77.25%	6.80%	-	-	-
1997	1272	2958	79.82%	5.14%	15.04%	30.87%	54.09%
1998	1417	3298	82.66%	4.24%	22.38%	38.93%	38.69%
1999	1600	4655	80.45%	5.11%	22.08%	37.79%	40.13%
2000	1774	7310	80.64%	5.43%	21.45%	41.07%	37.48%
2001	1996	16564	88.16%	9.51%	13.02%	45.97%	41.01%
2002	1976	20324	89.52%	4.32%	18.90%	53.53%	27.57%
2003	1987	21303	91.25%	3.12%	20.38%	49.15%	30.47%
2004	2046	23796	93.18%	3.69%	21.45%	48.63%	29.91%
2005	2001	20856	93.28%	3.17%	24.98%	45.57%	29.45%
2006	1805	19680	92.56%	1.57%	24.05%	46.61%	29.34%
2007	1819	21971	87.83%	1.76%	21.66%	44.06%	34.29%
2008	1849	23250	93.38%	2.06%	30.22%	40.17%	29.60%
2009	1765	20208	91.99%	2.07%	35.26%	43.65%	21.09%
2010	1475	15690	95.98%	1.72%	38.82%	38.77%	22.41%

 Table 1.3: Characteristics of options granted, 1996-2010.

Columns 6-8 give my estimates of the proportion of options that were granted at a similar time from year to year, which I term "scheduled" options. Their antipode are the proportion of options whose timing varies a lot from year to year, which I term unscheduled. In addition, it shows the proportion of options that I cannot assign to either category because I do not have consecutive year data. Since the coverage of firms increases over time and becomes more consistent, the percentage unclassified declines from a high in 1997 to smaller levels thereafter, albeit with year-to-year variation.

The key statistics are the proportions of options that I classify as scheduled versus unscheduled. Between 1997 and 2002 the proportion of options in the unscheduled group increased relative to the proportion in the scheduled group, while from 2003 to 2010 the proportion of scheduled options increased relative to the percentage of unscheduled options, so that in 2010 the proportion of options in the scheduled group becomes roughly the same as the proportion of options in the unscheduled group. In appendix 1.A I discuss other methods for differentiating scheduled and unscheduled options, which give similar results.

Since I estimate backdated options are those set at the minimum of a share price valley backdated options should be disproportionately unscheduled. Computing the percentage scheduled and unscheduled for options in the estimated backdating group, I find that 42% of backdated options are unscheduled while only 18% of backdated options are scheduled. On the contrary, 45% of non-backdated options are unscheduled compared to 25% of non-backdated options are scheduled.

Table 1.4 gives statistics on the level and change in another feature of executive stock options - the vesting period before executives can exercise the options. It shows a convergence of the vesting period of options during the fiscal year from 1996 to 2010. On the one side, "long term" options - those with a vesting period of more than five years virtually disappear. In 1996 these options constituted more than 24% of total transactions in 1996. In 2010 they made up less than 3% of transactions. On the other side, "short term" options, with 0-1 year time delays before they can be exercised also fall sharply from 21% in 1996 to 8% in 2010. One year vesting period options held steadily at 30% of total transactions. The big positive change is in the two and three year vesting period options, whose share of transactions increased from roughly 8% to 25% in 14 years.

The final column gives the maturity of the options - the period of time over which they may be exercised once they have been vested. Less than 1/4 of options have a maturity less than 10 years. On average, stock options have roughly a maturity of 9 years, but the vast

majority has a maturity of 10 years. Moreover, the proportion of options with a ten-yearmaturity has changed over time. In 1996 65% had a maturity of ten years, in 2004 this percentage increased to 84% in 2004 and then dropped back to 66% in 2010.

Year	Firms	Transaction	Vesting Period in Years					10 Years Maturity	
			0	1	2	3	4	≥5	
1996	728	1455	21.17%	34.16%	7.70%	8.11%	4.33%	24.54%	64.67%
1997	1272	2958	25.32%	33.30%	9.20%	8.01%	3.65%	20.52%	66.29%
1998	1417	3298	24.38%	32.38%	9.25%	8.73%	5.12%	20.13%	67.43%
1999	1600	4655	16.15%	28.38%	14.48%	13.77%	8.06%	19.16%	74.91%
2000	1774	7310	11.53%	23.87%	19.66%	18.48%	10.74%	15.72%	76.83%
2001	1996	16564	9.22%	23.35%	23.86%	21.15%	12.56%	9.86%	83.76%
2002	1976	20324	10.24%	24.12%	24.87%	22.18%	12.41%	6.17%	83.02%
2003	1987	21303	10.28%	23.91%	25.63%	22.98%	12.64%	4.56%	82.32%
2004	2046	23796	8.52%	23.36%	25.99%	23.76%	13.32%	5.05%	83.97%
2005	2001	20856	9.03%	23.83%	26.29%	24.48%	12.53%	3.83%	78.50%
2006	1805	19680	8.33%	23.58%	26.27%	24.76%	12.94%	4.12%	71.06%
2007	1819	21971	12.70%	22.76%	25.39%	23.44%	11.99%	3.72%	65.91%
2008	1849	23250	8.06%	23.10%	26.87%	25.42%	12.92%	3.63%	69.29%
2009	1765	20208	8.51%	24.59%	26.84%	24.86%	12.03%	3.18%	63.13%
2010	1475	15690	8.01%	24.75%	27.32%	25.21%	12.07%	2.65%	66.18%

Table 1.4: Stock options vesting period and maturity.

On average a firm grants options to CEOs twice a year (1.99 days), with a standard deviation of 5.33 days. In table 1.5 I report how many distinct days in a year a firm grant options to CEO. It can be inferred that more than 50% of transaction occur in a single day.

#### 4. Linking Execucomp and IFDF

Execucomp and IFDF differ in the information on stock-based compensation in important ways that justifies my use of examining both datasets.

The advantage of the IFDF is that it contains detailed information on derivative transactions by capturing all of the insider activity reported in SEC forms. The detail of the IFDF data makes it invaluable to any analysis of options. The IFDF reports the number and

type of derivatives granted, the transaction date, the date the SEC received the file, when it was signed by the officers, the name and the role of the officers granted the derivatives (or who exercised the derivatives), the strike price, the maturity date, the exercise date, the number of shares underlining each derivative (as well as the number of share adjusted for an eventual stock split), and the name of the person and firm in the transaction. It also evaluates of the accuracy of the data reported by the insiders, including in some cases its own estimate of the transaction date and an indicator that reflects its views of the accuracy of the data. It corrects the data that it judges inaccurate. All this information is indispensable to determining if the options were granted at the money, out of the money or in the money, and whether the options were scheduled, unscheduled or backdated. Finally, the IFDF provides information on exercised options, forfeited options or other dispositions of options. It is a massive body of data that illuminates the "world of the derivatives/options".

		2	3	4	>5
year	1 day	days	days	days	days
1996	611	192	45	4	57
1997	983	464	111	40	109
1998	1093	498	165	56	57
1999	1266	506	165	60	72
2000	1326	672	228	96	84
2001	1510	710	243	112	146
2002	1553	646	204	88	91
2003	1610	588	144	100	54
2004	1710	534	117	80	160
2005	1664	538	117	84	70
2006	1545	434	93	32	28
2007	1572	394	87	64	32
2008	1576	438	129	36	17
2009	1542	370	69	40	34
2010	1309	290	45	20	11

Table 1.5: Granting days per year.

Execucomp reports all type of compensation at fiscal year frequency (Kuhnen *et al.*, 2012) for top executives in public US companies, including all S&P 1500 companies, which constitutes approximately 80% of the total market capitalization in U.S. (Bebchuk *et al.*, 2005).

As we have seen this information includes salary, stock options awarded, bonuses, long-term incentive plan, and restricted stock. The information on stock options is not as detailed as that reported IFDF and, as noted, changed when FAS123R changed the reporting rules in 2006, creating a problem of time series inconsistency.

To use both data sets in concert, I have merged them together. This not an easy task due to the different ways the data are constructed and reported. The IFDF collects information on every transaction in a company involving insider's stock options and stocks from Form 3,4,5, and 144. This means that for the same company in a calendar year, I have as many observations as the transactions occurred for each executive. By contrast, Execucomp data are for fiscal years. Thus, one has to amalgamate the IFDF data into a yearly basis for any match.

Both data sets give the names of the officers receiving options but they can give differently the insiders' name (order, abbreviation, middle name), so that an exact match of names loses observations due to the differences between Charles M. Jones and C.M. Jones, and so on. The name disambiguation problem is a well-studied one in the bibliometric literature, where computer scientists have written extensive codes to match names, based not only on the name but also on the address and even on the past history of the person. Examples of recent works that have merged Execucomp with Thomson through officers name are Knewtson (2011) and Ladika (2012).

For the match, I transform the IFDF data into an annual file for CEOs that I can match to the Execucomp data on CEOs. I exclude all insiders' activities in the IFDF relating to executives who were not a CEO. I focus on officers that are indicated as CEO by Execucomp, which has a flag for CEO if the officer served as CEO for most of the year. Since Table II reports all type of derivatives granted to top officers, I define stock options the following derivatives: Call Option, Options, Non-Qualify Stock Option, Employee Stock Option, Director's Stock Option and Non-Employee Director Stock Option. I then sum up the number of derivatives reported in IFDF by company using the company's PERMNO and fiscal year and merge the data with Execucomp. I am able to merge 44% of the information regarding CEO provided by IFDF with Execucomp.

There is one piece of information in the two data sets that provide me with a potential measure of the consistency of the data. This is the number of options granted. Both IFDF and Execucomp report this statistic. If my matches were perfect and the firms reported the same data on the insider trading file as on the corporate shareholder statements, and there were no glitches in the reported data, then the number of options granted would be the same in the two data sets.

But there are potential problems that can create a divergence between them. First, Execucomp identifies the officer that has the function of CEO for most of the fiscal year. Therefore, I may under-report information provided by Execucomp if a CEO changed during the year. Second, in 75 cases the IFDF reported that CEOs were granted options under a different main role than CEO. For example, an officer reported he was granted options as a Director under "rolecode1" which is the main role, as a Chairman of the Board as a second role (rolecode2), as President as rolecode3 and finally as a CEO as rolecode4. For those 75 cases the Execucomp data and the IFDF data would differ in the number of options a CEO was awarded. Third, before 2002, SEC allowed officers to report the transaction within 40 days of the purported date at which it was made.<sup>6</sup> This means that some options on the IFDF file might be assigned to a wrong fiscal year. For instance, a transaction occurred in the months of December might be assigned to the month of January, and thus to a different fiscal year if the fiscal year coincide with calendar year, as is often the case.

These problems notwithstanding the number of derivatives reported by IFDF are strongly correlated with the number of derivatives reported by Execucomp in most years.

<sup>&</sup>lt;sup>6</sup> "Until August 2002, the requirement had only been to file Form 4 with the SEC within ten days after the close of the calendar month in which the transaction had occurred" (Brochet, 2010, p. 420).

Figure 1.3 below shows the correlation between the number of derivatives reported in Execucomp and the ones reported in Thomson.

The two bars in the figure show the number (in million) of total options granted by all company to their CEOs as reported by Thomson Inside File Form, Table II and the options reported by Execucomp for CEOs with the same name in each fiscal year. As can be seen the number of options is highly related and moves together over time. The correlation between the two measures for all years and executives is 0.91. But this correlation masks considerable heterogeneity from year to year in the tightness of the link between the IFDF and Execucomp reports.



Figure 1.3: Total number of stock options granted to CEOs reported by IFDF and Execucomp by fiscal year.

Table 1.6 below shows correlations that range from 1 in 2009<sup>7</sup> to 0.43 in 2006. The low correlation in 2006 presumably reflects the FAS123R changes in the reporting rules of the DEF14A form. Most of the correlations within each year are, moreover, lower than 0.91 correlation for the entire period. This reflects the reasonably tight relation between the overall levels of the options granted in both data sets over time, as displayed in figure 1.3.

<sup>&</sup>lt;sup>7</sup> The high correlation explains why the total number of stock options granted to CEOs reported by IFDF and Execucomp by fiscal year reaches a maximum in 2009.

In sum, there are inconsistencies between the number of options reported in the two data sets, but the correlations are sufficiently high except for 2006 when Execucomp changed their mode of reporting that combining the two data sets should give us a more complete and picture of the pattern of executive compensation over time than examination of each data set by itself.

Table1.6:betweentheoptionsgranestimatedfromreported by E	Correlation number of ted to CEOs m IFDF and xecucomp.
Year	Corr.
1997	0.86
1998	0.83
1999	0.65
2000	0.87
2001	0.74
2002	0.88
2003	0.88
2004	0.83
2005	0.81
2006	0.43
2007	0.88
2008	0.74
2009	1
2010	0.78

#### 5. Conclusion

This chapter has explored the changes in US CEO compensation from 1996 to 2010 using data from the Execucomp dataset and the IFDF dataset. It has shown that CEO pay rose by \$2.7 millions between 1996 and 2010, but the changes varied greatly over time with the state of the stock market and economy. Most of the increase in stock-based compensation during the period from 1996 to 2001 was due to increased grants of stock options, but options became less popular toward the end of the period in favor of direct grants of stocks, which in

2010 constituted almost 35% of total compensation. The chapter has also shown substantial changes in the composition of options over time. The percentage of scheduled options increased by almost 25% between 1997 and 2010. Backdated options disappeared.

Cash-based compensation increased more modestly over the period. The chapter described the methodology by which I matched Execucomp data from company's annual proxy statement (DEF 14A SEC) with data from the Insider Filing Data Feed on U.S. insider activities. I found that the number of options reported in the two datasets is highly correlated (except in 2006 when Execucomp changed the mode of reporting options) to allow me to use them together to analyze some of the changing patterns in executive compensation, on which the rest of this thesis focuses.

#### **Appendix 1.A: Some details on constructing the dataset**

To construct my data set I use three different datasets. Executive compensation data are taken from Thomson Financial Insider Filling (IFDF) database and Compustat. I obtain detailed information of firms' daily stock price from Center for Research in Stock Price (CRPS).

#### **1.A.1** Thomson Financial Insider Filling Database

From IFDF Table II I obtain the official grant date, the exercise date, the data at which options expire and the exercise price of insiders' grants. I restrict my analysis on transactions occurred under transaction code A during the period from 1996 to 2011. Transaction code A reports award transaction pursuant to Rule 16b-3(C).

Table 1.7 gives the details of how I obtained the data for my analysis beginning with 4,006,678 transactions reported in the IFDF file. I use yearly official company's Ticker and company's CUSIP number to match companies from IFDF database to Center for Research in Stock Prices databases (CRSP). I drop companies that were not matched with CRSP. I therefore lose 931,359 transactions. CRSP allows me to assign the PERMNO (permanent security identification, that is unchanged during the time) for each firm and the closing price for each transaction. I also exclude non-option derivatives (573,065), so that I consider Call Options, Options, Non-Qualifying Stock Options, Employee Stock Options, Director's Stock Options and Non-Employee Director Stock Options. I first retain information for top five executives - CEO, Chairman of the Board, Chief Operating Officer, General Counsel and President). This reduces my sample by 2,018,469 transactions.

I lose a small number of transactions due to what appear to be data errors in the sample. I found some observations where the strike price was not reported by IFDF (4,961); other observations which gave an exercise price before the grant date (952); others which misreported the maturity (6,218) date or reported the maturity date before the transaction occurred (52); and one observation reporting the number of derivatives granted less than zero. This left me with 471,601 grants. I use CRSP/Compustat Merged dataset to infer for each permanent security identification number (PERMNO) the corresponding variable FYR (indicating the months on which the fiscal year ends). I am thus able to link the CRSP permanent security identification number (PERMNO) with the variable FYR provided by Compustat and allocate the options award reported in Thomson to the right fiscal years from 1996 to 2010. For instance, an option granted in February 2011 would be assigned to fiscal year 2010, if the fiscal year ends in May 2011. For 13,882 (2.94%) transactions I am not able to find information about FYR, I thus assume that fiscal year (Execucomp). I drop transactions assigned to 2011 fiscal year (34,627) or fiscal year 1995 (73). I end up with 436,901 transaction occurred between fiscal year 1996 and 2010. Eventually, I restrict my analysis on CEOs only, obtaining 223,318 transactions.

IFDF	Transaction
Transaction A	4006678
No PERMNO information	-931359
Derivatives different from Options	-573065
Non top five executive	-2018469
Top five executive Options Space	483785
Misreported Transactions	-12184
Occurred in Fiscal year>2010	-34627
Occurred in Fiscal Year<1996	-73
Non-CEOs	-213583
CEOs	223318

Table 1.7: The accounting procedure. The sample extractedfrom IFDF.

To get the grant data for my analysis I then proceeded as follow:

- If the reported exercise price is equal to stock closing price inferred from CRSP, the grant date is equal to the reported one. If not, I checked the previous two days and selected the grant date that minimized the difference in absolute value between the stock closing price and the strike price. I defined an option at the money if the exercise price exactly matched or was within 10% of the closing price listed in CRSP. I estimated that approximately that 90.72% options are granted at the money, 5.67% is granted out of money and the remaining are granted in the money.
- I then obtained daily stock price over a window of 41 days (the inferred grant date plus 20 trading days before and after). I defined an option as backdated if the stock closing price on grant date is the minimum over the 41 days window. I defined an option as scheduled if it was awarded inside of 2 weeks window of the awarding of options in the previous year.<sup>8</sup> That is, for any company I look if it grants an option one week before or one week after one-year anniversary to a previous transaction. I defined an option as unscheduled if it was granted in period of time outside of the 2 weeks window around the reported transaction date in the previous year. If no options were awarded during the prior year I defined the option as "unclassified". I followed a similar procedure proposed by Lie (2005) in deciding which options are scheduled. I experimented with different windows and obtained results comparable to those described next.

To check the robustness of my assignment of options as scheduled and unscheduled, I experimented with different time windows to see if the choice of window-days affects the percentage the esteem of scheduled options. I simulated five scenarios with different days window for identifying scheduled options. In the first scenario, I define an option as scheduled if it was granted within a window of fourteen days - seven days before and seven days after- of

<sup>&</sup>lt;sup>8</sup> Some options are considered scheduled even if no options were granted the previous calendar year. For example, a company granted no options in 1999. In 2000 the firm granted some options at the begin of January, and some others at the end of December. If the letter ones were granted in a window of two weeks from the award made in January, I consider them as scheduled.

one-year anniversary of the previous grant or within fourteen days of a next-year option grant. The only difference between this definition and the definition in the body of the chapter is that I consider options as scheduled if they were granted for the first time but were then granted in the same period in the next year. The second scenario considers options as scheduled if they are granted within twenty days of one-year anniversary of the previous grant. The third and fourth scenarios consider options as scheduled if they were granted if they were granted sixteen or twelve days respectively of one-year anniversary of previous grants. The last scenario considers options as scheduled if they were granted in the same month of one-year anniversary of the previous grant. I summarize the number of scheduled transactions according to different scenarios in table 1.8 below.

The data shows similar trend to that found in the seven days window used in the chapter.

Fiscal	Total	Scenario 1:	Scenario 2:	Scenario 3:	Scenario 4:	Scenario 5:
Year	Transactions	14 days	20 days	16 days	12 days	same month
		window	window	window	window	
1996	1466	352	-	-	-	-
1997	2996	997	523	458	423	558
1998	3363	1177	799	752	715	845
1999	4744	1468	1114	1052	960	1213
2000	7371	2349	1698	1607	1508	1637
2001	16877	4237	2599	2222	2036	2826
2002	20607	6123	4479	3952	3716	4856
2003	21423	6979	4977	4567	4185	5225
2004	23893	7938	5982	5235	4840	7127
2005	21211	8236	5692	5396	4858	6639
2006	19804	7431	5209	4842	4524	5577
2007	21902	8743	5537	4943	4535	6010
2008	23050	10253	7656	7214	6706	8154
2009	20290	9570	7885	7223	9227	8368
2010	15397	8293	6603	6231	5745	7096

 Table 1.8: Scheduled options according to different scenarios.

#### 1.A.2 Compustat

I collected detailed information on the amount and composition of top manager compensation from Standard & Poor's Executive Compensation data set (Execucomp) for fiscal year 1996 through fiscal year 2010. Execucomp collects top executive compensation detailed information directly from company's annual proxy statement (DEF 14A SEC). Most companies report data on only five executives but some companies voluntarily report more executives. I conducted my analysis on CEOs only. The variable CEOANN provided by Execucomp indicates which officer served as CEO for the entire or most of the fiscal year. In 20 cases Execucomp indicates that two officers served as CEOs for the same firm. I therefore used the mean of the data on the two officers as the CEO statistics.

In 2006 the FAS123R changed the reporting rules. Under the new rules, companies have to report the estimated fair value of the stock options granted in the proxy statement. Execucomp provides its own fair award value estimate for stock options granted before 2006 using a non-standard Black-Scholes formula. In 2006, Execucomp also changed the way total compensation is calculated. Before 2006, total compensation (TDC1) was calculated as the sum of salary, bonus, restricted stock granted, stock options, long-term incentive payout and other forms of compensation. After 2006, total compensation was calculated as the sum of salary, bonus, stock awarded, stock options, non-equity incentive plan compensation, deferred compensation earnings reported as compensation and other compensation.

I obtained from Execucomp the following information: the dollar value of salary (Execucomp variable SALARY in thousands of dollars), bonus (Execucomp variable BONUS in thousands of dollars), the value of stock options awarded (Execucomp variable BLK\_VALUE until 2006, rename as AWARD\_FV thereafter), the value of stock awards (Execucomp variable RSTKGRANT before 2006, renamed STOCK\_AWARD\_FV thereafter), the Long Term Incentive Payout (LTIP), Deferred Compensation Earnings Reported as Compensation (DEFER\_RPT\_AS\_COMP\_TOT, available only after 2006) and All Other Total (ALLOTHTOT substituted in 2006 by OTHCOMP). I also collect the number of derivatives granted to each manager (Execucomp variable OPTION\_AWARD\_NUM).
As pointed out in the text, in 2006 FAS123R changed the reporting rules. Specifically, after 2006 companies have to report the estimated fair value of the stock options granted in the proxy statement. Different evaluating methods are permitted, including the Black-Scholes and the binomial options pricing model. Before 2006, Execucomp provides its own fair award value estimate for stock options granted using a non-standard Black-Scholes formula. Therefore, the time series include the value of stock options granted estimated by Compustat from 1996 to 2005, and the value of stock options granted evaluated by each company since 2006.

Chapter Two: Price Targets Accuracy and Timing: Evidence from IBES's Price Targets Data Set

# 1. Introduction

This chapter describes the Institutional Broker Estimate System (IBES) Detailed History Price Target data file that provides statistics on analysts' price targets - "the projected price level forecast by an analyst within a specific time horizon" (Glushkov, 2009, p. 6) - that I use to analyze the arms-length bargaining and managerial power models in the rest of this thesis. Because price target data are not analyzed extensively in the financial forecast literature and have never to my knowledge been used in analyzing executive compensation, I provide a brief discussion of the variable and existing knowledge of the price target process in section 2. I give a detailed description of the IBES price target data in section 3. Section 4 analyzes three issues regarding price targets that are critical to my ensuing analysis of the relation between price targets and executive compensation: the scheduling of price target announcements over the year; the extent to which the target prices are positively associated with share prices at the forecast date; and the extent to which the market treats announcements as new information that gets embodied quickly into share prices.

# 2. Background

Stock market analysts give price target forecasts for the share prices of firms several times a year, so that the price targets provide a flow of information about experts' views of the future performance of companies that market participants can use in investment decisions and the management and boards can use in setting executive pay.<sup>9</sup> But "Compared to EPS<sup>10</sup> forecast accuracy studies, the literature on target price is much more recent and substantially less populated" (Bilinski *et al.*, 2011). Of the approximately 250 papers on financial analyst forecasting in Ramnath *et al.*'s 2008 literature review of the top eleven research journals since 1992, I counted only four papers containing the word "price target(s)" or "target price(s)" in the author's key points summary, and only two papers with such words in the article bibliography.<sup>11</sup> A likely reason why analysts have not examined price target data is that the IBES price target data set was not available to academic subscribers until 2009 (Bradshaw *et al.*, 2012).

Asquith *et al.* (2005) gives summary statistics on financial analyst reports in 1997-1999 using data from sources of IBES competitors. The statistics show that analyst reports typically include buy or sell recommendations and earning forecasts, but contain price targets only 72.6% of the time<sup>12</sup>, and that reason is that analysts tend to avoid issuing price forecast for unfavorable reports. About 95% of recommendations to hold or buy include price targets while only 65.8% of recommendations to sell include price targets. Relatively few analysts give price targets below the share price. They suggest that analysts might use forecasts of price targets to keep positive relations with the managers on whom they rely for information.

Clement (1999) examines the IBES Detailed History tape for the period 1983 to 1994, such as the number of analysts, the number of forecasts, the number of brokers and the number of firms, but could not price targets, which were not available until 1999. Bilinski *et al.* (2011)

<sup>&</sup>lt;sup>9</sup> To prelude, I argue in chapters 3 and 4 that the target price should affect the composition of executive composition differently if a board sets pay by arms-length bargaining than if CEOs greatly influence the paysetting process. When the target price is high, an option is worth more to the executive so that an arms-length board could give the executives smaller options or adjust the exercise price for the likely change in value that is independent of the executive doing anything out of the ordinary. By contrast an executive-dominated board would give bigger options to the executive as a means of enriching that persons.

<sup>&</sup>lt;sup>10</sup> EPS stands for earning per share.

<sup>&</sup>lt;sup>11</sup> Three papers use Investext, which stores research reports written by analysts and is provided by Thomson; two use First Call, one of the main competitors along with Zacks, in the population of analysts covered, the variables reported and the time series (Glushkov, 2007) and one also uses Real Time Database. In 2001 Thomson Reuters acquired First Call, merged it with IBES, and then discontinued it (Bilinski *et al.*, 2011).

<sup>&</sup>lt;sup>12</sup> Brav *et al.* (2003) also finds that only 2/3rds of all analyst reports includes price targets.

studied the variation of price target accuracy across 16 countries in the IBES Detailed History data for the period between 2002 and 2009. They provide detailed descriptive statistics on the distribution of target prices, firms, brokerage houses, analysts and price target accuracy by country. The authors found that the US data, which constitutes 55.2% of the entire price targets sample and 44.8% of the sample firms, dominate the IBES price target data. Australia and Hong Kong offer the highest proportion of met<sup>13</sup> price targets, with 66.1% and 64.3% of price target met respectively, whereas the US percentage is only 54.7%. The differences are mainly due to accounting quality disclosure, corporate governance, culture traits and regulations. Comparing the forecasts of analysts, Bilinski et al. (2011) finds that some analysts have superior forecasting ability and rejects the notion that price targets are "just for show" as the popular press claims. Cowen et al. (2006) argue that optimistic reports on dot-com stocks before the dot.com collapse and the Enron scandal have damaged the reputation of analysts: "popular explanation for all of those failures is that analysts working for investment banks were compromised by the hefty bonuses they could earn writing positive reports on investment banking client, or were pressured to write favorable reports by investment bankers at their firms" (Cowen et al., 2006, pp.120). Cowen et al. (2006) examine differences in analysts' optimism in recommendations and forecasts (including price targets) by analysts working for full-service banks (which financed the research through brokerage and underwriting new issues); syndicate banks (which sustained the research through fees and trading), brokerage firms (that financed the research through trade commission and rewarded analysts in base of the trading volume of the stock) and pure research firms (that sell research itself). They show that full-service banks employed the highest number of analysts who in turn covered the biggest number of firms and found that bank analysts issued the less optimistic price target forecasts.

<sup>&</sup>lt;sup>13</sup> The authors use two measures to capture accuracy. The first one is a dummy variable taking the value of the unity if the share price reaches or overcome the price target any time over the 12 months forecast horizons. The second one measure the forecast error, and it is designed as the difference in absolute value between the share price at the end of the 12-months period and the price target divided by price at the forecasted issue date.

Of particular relevance to executive compensation, Francis *et al.* (1993) and Lim (2001) show that company's managers can pressure analysts to modify reports by exploiting the analysts' dependence on managers for information. In extremum circumstances, firms have fired analysts writing bad reports about clients on whom the firm depends (Richardson *et al*, 2004), like Chung Wu, UBS analyst, fired under pressure from Enron for having advised clients to disinvest from Enron on the night of the August 21, 2002 (Oppel, 2002). Richardson *et al.* (2004) show that firms' managers manipulate analysts' reports to have optimistic earning forecasts and to subsequently correct the forecast so they can surpass it. By doing so, managers can sell the stock on behalf of the firms or for personal accounts after having announced that they actually had beaten the target and inflating the stock price. McAnally *et al.* (2008) show that managers also game the system in the opposite direction, missing targets on purpose before options grants to benefit from a lower strike price. These forms of behavior play a part in my chapter 4' analysis.

# 3. Data and analysis

The data I use on analysts price targets comes from IBES Detail Adjusted History Price Target data file for the period from calendar year 1999 to 2011. The IBES file provides summary measures of the targets from all analysts who make forecasts in a given period and also gives the underlying price targets of individual analysts. I use the summary measures to indicate the likely direction of future share prices, which will determine the expected value of stock options. Given that my CEO compensation data are for US firms, I focus on US firms only, though the IBES includes price targets for firm around the world.<sup>14</sup> I use the adjusted file

<sup>&</sup>lt;sup>14</sup> Thomson-Reuters reports in its discussion of the IBES data: "the reporting currency does not reflect the clear majority of estimate submissions, Thomson Reuters may exercise the option to set the default based on the currency of the majority of estimate submissions. In cases where companies report in multiple currencies,

data set (which IBES calls as Normalized) that adjusts value for various corporate actions, like splits of stocks and that dividend, and that is comparable across currencies.<sup>15</sup>

The IBES Price Target data file records the following variables: analysts' price targets level, analysts' name, the company he/she works for, company for which he/she issues the target price, the horizon-period, the day the price target was announced and when became active in IBES data file, the company currency and whether or not the company is a US firm.<sup>16</sup> The price targets are available March 1999 to the present.

Table 2.1: Analysts' horizons.						
Horizon	Observations	Percentage				
0	21	0.002%				
1	586	0.068%				
3	413	0.048%				
5	3	0.000%				
6	11891	1.388%				
7	2	0.000%				
8	1	0.000%				
9	423	0.049%				
10	2	0.000%				
11	1	0.000%				
12	837905	97.808%				
13	1	0.000%				
14	1	0.000%				
15	14	0.002%				
18	3868	0.452%				
24	962	0.112%				
36	590	0.069%				
50	1	0.000%				
Total	856685	100%				

Thomson Reuters will set the default currency based on the majority of estimate submissions." (I/B/E/S Detail History-User Guide, 2009, p.11).

<sup>&</sup>lt;sup>15</sup> IBES distinguish US firms from international through a dummy variable (USFIRM) that take the value of zero if the firms analyzed is from international file and one if from US file. Canadian firms however might report under both flag. IBES puts an "@" symbol in front of the tickers of international company (with the exception of Canadian). One way to obtain Canadian firms is to filtering USFIRM=0 and look for company whom ticker does not have a "@" in it. However, a different issue is to separate Canadian companies from US file. Since I merge IBES with CRSP, which provides institutional details on the company, the problem disappears.

<sup>&</sup>lt;sup>16</sup> Despite limiting my data to US firms, 983 observations report that the currency at company level is not the USD, while 1,649 observations report that the estimated currency is different from the US dollar. In 358 cases both currency at company level and estimate currency are not in US dollars. However, since the data are taken from price target adjusted file, values should be comparable across various currencies and expressed in USD. I clarified these issues by e-mail with WRDS support.

The targets cover different horizons expressed in months, but as Table 2.1 shows 98% of price targets have a 12-month horizon, so I limit my analysis to those targets.

I use CUSIP and TICKER provided by IBES to infer the PERMNO code on Compustat and CRSP. I drop 92,090 observations for which was not possible assign a PERMNO.

For every year I count the number of analysts, the firms they work for, the companies they cover and the number of price targets they issue. In 18.65% of the cases IBES reports more than one announcement in one day. From CRSP I obtain the share price exactly one year after the 12-month price target forecast. Among 19,709,601 closing prices reported by CRSP for the IBES companies, I track 12,879.517 closing prices exactly one year later. However, some closing prices are issued on Friday or before holidays so I do not have a price exactly one year later. In these cases, which make up 4,040,619 observations, I take as effective price, the price reported one year and three days later. I next define a price target as being accurate if the share price matches or exceeds the effective forecast price if an increase in share price is forecast) or if the share price falls below the effective share price when the share price is expected to fall.

	No.	No.	No.	No. price	%		%
	brokers	analysts	firms	targets	positive	accuracy	accuracy
1999	156	2381	3231	21865	0.61	10703	0.49
2000	197	2933	3475	31527	0.66	17370	0.55
2001	167	3346	3213	37997	0.69	22949	0.6
2002	180	3271	3173	44788	0.68	24196	0.54
2003	253	2823	3292	48471	0.58	14728	0.3
2004	290	2906	3586	53131	0.62	22684	0.43
2005	295	2937	3789	54823	0.69	25226	0.46
2006	284	2914	3913	58490	0.74	28905	0.49
2007	278	2987	4028	63993	0.78	46753	0.73
2008	281	2962	3836	76863	0.82	54430	0.71
2009	302	2887	3600	77841	0.74	27350	0.35
2010	331	3289	3756	82814	0.81	40653	0.49
2011	316	3412	3829	93212	0.84	-	-

Table 2.2: Summary statistics by years.

Table 2.2 gives the summary statistics of the price target forecasts in my data set.

Columns 1-3 show increases in the number of brokers, analysts, and firms that make price targets. Column 4 shows a 4.26 increase from 1999 to 2010 in the number of price targets in the data set. Column 5 shows that majority of price targets expect an increase in the share price. In doing this analysis, I encountered some price targets that seemed implausible. For instance, for TROY GROUP INC, IBES reported a price target of \$4,200,000 in 06/April/2000, when the price was \$22.56.17 Some of these data problems reflect the CRSP stock closing price. CRSP reports for Berkshire Hathaway a share price of \$2,411 in 02/01/2001, which ultimately grew up to \$3,476 in 20/01/2010. The day after, i.e. 21/01/2010, the share price reported by CRSP was \$72.72. A closer look to the dataset reveals that the firm split the shares at 50 to 1 on January 21. Analysts presumably issued a price target basing the forecast on the share price without knowledge of an eventual split. Firms buying back share presents a different problem, as analysts might try to estimate the likelihood of such an event. If analysts do not take into account a buyback, they will issue what would look like an extremely pessimistic and inaccurate forecast. For 40 companies, 48 price targets were for a price of zero. When I queried Thomson-Reuters WRDS support about extremely large or small targets, they suggested that the most likely explanation of weird observations are data errors on part of Thomson.<sup>18</sup> But to be conservative, rather than defining some price target as implausible ex ante, I keep all observations.<sup>19</sup>

Table 2.3 divides some of the statistics in table 2.2 to show the work load of firms and analysts in producing price targets. A typical broker firm reports price target for 81.98 companies with a standard deviation of 171.19 companies.<sup>20</sup> Table 2.3 reveals that, on average

<sup>&</sup>lt;sup>17</sup> For 376 (0.05%) observations, the difference between the price target and the share price at the day of the announcements was 1,000 or more. Similarly, for 44 firms 360 (0.048%) price targets foresee a change in price of 10,000% or more.

<sup>&</sup>lt;sup>18</sup> As suggested by e-mail by the WRDS support.

<sup>&</sup>lt;sup>19</sup> I also conduct the analysis dropping those observations and I find that the results do not change or they get slightly better.

<sup>&</sup>lt;sup>20</sup> The percentage of broker firms that next year cover less companies is 55.10%, while the percentage of them that cover ten or more additional firms is 30%. To calculate the percentage I construct a panel, which reports the total number of different company covered by each broker firm in a year. I then look how many broker companies in a year experienced a decrease in the number of firms they covered in a year.

a broker firm employs 13.79 analysts and announces 223.97 price targets per year while an analyst covers 7.17 companies per year and makes 19.10 price targets per year.

	Avg. announcements per Broker firm	Avg. analysts per broker	Avg. companies covered per broker	Avg. companies covered by single analyst	Avg. analysts serving broker firm	Avg. price targets per analyst
1999	140.16	19.23	78.62	5.13	1.26	9.18
2000	160.04	20.14	78.33	5.2	1.35	10.75
2001	227.53	28.99	100.06	5.13	1.44	11.36
2002	248.82	23.73	99.34	5.7	1.31	13.69
2003	191.59	12.75	68.79	6.39	1.14	17.17
2004	183.21	11.26	67.76	6.94	1.12	18.28
2005	185.84	11.09	70.4	7.25	1.11	18.67
2006	205.95	11.05	76.82	7.75	1.08	20.07
2007	230.19	11.48	82.91	8.08	1.07	21.42
2008	273.53	11.37	83.84	8.24	1.08	25.95
2009	257.75	10.22	81.51	8.7	1.07	26.96
2010	250.19	11.1	88.14	8.96	1.12	25.18
2011	294.94	12.03	97.03	9.23	1.11	27.32
Tot:	223.97	13.79	81.98	7.17	1.18	19.10
Std.Dev	541.86	26.13	171.19	6.63	0.61	23.77

Table 2.3: Typical broker firm and analyst.

Table 2.4 records the statistics on price targets in terms of the number of companies covered by brokers. The first row indicates that 8,917 companies are covered by a single broker firm for a given year, and that on average one analyst from the firm issued almost two price targets for that company. Similarly, the second row indicates that 6,467 companies are monitored by two broker firms which have about 2 analysts who announce 4 price targets per year, and so on.

The first two columns show a power law relation with many companies covered by small numbers of brokers and a few companies covered by many brokers. The majority of firms are covered by 4 of fewer brokers while 456 firms are covered by 19 or 20 brokers. For example, the large corporation, Oracle Corporation, was covered in 1999 by 19 different brokers who reported 56 price targets.

~		Avg. Analyst		
Company	Fraguanay	per Broker	Avg.	Avg.
	8017	1.04	1.07	
1	6467	2.00	1.97	2.06
2	5121	2.09	4.21	2.00
3	JIJI 4418	3.15	0.7	5.06
4	4418	4.2	9.29	4.11
5	3649	5.20	12.30	5.12
6	2970	0.31	13.10	0.14
/	2423	7.4	18.26	/.18
8	2147	8.47	21.58	8.2
9	1705	9.51	24.61	9.24
10	1373	10.6	27.95	10.24
11	1243	11.7	31.74	11.33
12	1027	12.73	35.2	12.29
13	856	13.85	38.04	13.35
14	736	14.86	41.87	14.32
15	637	15.92	45.28	15.33
16	480	16.97	49.42	16.31
17	468	18.03	53.37	17.35
18	401	19.12	56.02	18.36
19	292	20.27	60.1	19.46
20	248	21.45	63.46	20.6
21	208	22.46	68.59	21.57
22	197	23.71	70.89	22.65
23	128	24.53	73.67	23.57
24	116	25.8	75.89	24.76
25	101	26.97	77.88	25.68
26	72	27.68	84.44	26.39
27	59	29.07	90.51	27.69
28	55	30.36	94.13	29.09
29	37	30.73	89.54	29.43
30	29	32.17	108.34	30.9
31	33	33.64	98.7	32.42
32	22	34.95	107.14	33.45
33	14	35.64	112.57	33.79
34	15	36.8	102.07	34.87
35	7	37.14	118.57	35.57
36	7	38	124.57	36.86
37	3	38	117	36.67
38	7	40.85	133.86	37.86
39	7	42.86	135.29	41.14
40	3	43.33	12867	40.67
41	6	43.83	130	41
43	1	46	108	40
44	- 1	48	152	43
45	2	48.5	143	45

Table 2.4: Typical analyzed company.

Table 2.4 (continued).						
46	1	50	156	48		
53	2	59	213.5	55.5		

The 19 broker firms had 21 analysts following Oracle, which implies that some firms have more than one analyst following the company.<sup>21</sup> The last column indicates the average number of analysts analyzing the firms.<sup>22</sup> Finally, as the "coverage" of a company (defined as the number of broker companies that analyses the company) rises, the number of announcements increases as well but more than proportionally.

Table 2.5 shows that the number of price targets given per company have risen, the result of both an increase in the number of brokers and firms and also a tendency for analysts to issue price targets more frequently (and possibly for IBES to have become more inclusive).

statistics.						
Year	Mean	Std.	Median			
1999	6.77	7.48	4			
2000	9.07	10.91	5			
2001	11.83	13.64	7			
2002	14.12	16.03	8			
2003	14.72	16.03	9			
2004	14.82	16.45	9			
2005	14.47	15.46	9			
2006	14.95	15.51	10			
2007	15.89	16.36	10			
2008	20.77	21.77	12			
2009	21.62	23.2	13			
2010	22.05	23.54	14			
2011	24.34	26.7	14			

Table 2.5: Announcements per company and year: summary statistics.

Figure 2.1 shows the distribution of announcements per company and calendar year.

<sup>&</sup>lt;sup>21</sup> A closer look reveals that the firm Firstalb and Piper had two analysts covering Oracle Corporation in 1999.

<sup>&</sup>lt;sup>22</sup> The number of analysts covering a firm can differ from the number of analyst per firms in the third column because an analyst may switch brokers in a year. For instance, for Noven Pharmaceut two different broker firms report a price target, one in November and one in May but the analyst that conducts the analysis was the same. But analysts do not change company often: on average an analyst serves 1.18 companies with a standard deviation of only 0.61 companies as reported in table 2.3.



Figure 2.1: Announcements per company and calendar year.

Why do analysts announce more target prices in some years than in others? One likely factor in determining the number of target prices in a year is the volatility of the share price. The more volatile the price, the more reason to update price targets. To see if this expectation holds in the data, I regressed the number of announcements per firms in a year on stock price standard deviation obtained from CRSP and company dummies. The results in table 2.6 below show that analysts issue updated price targets more often when the share price volatility was higher. The likely explanation is that the more volatile the share price is, the more difficult is to forecast a price.

Table 2.6: Multivariate analysis of the effects of share price volatility on the number of announcements per firms in IBES datasets from 1999 to 2010. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%, 5% and 1% levels respectively. The t-statistic are reported in parentheses.

	(1)	(2)	(3)
	ANNOUNCEMENT	ANNOUNCEMENT	ANNOUNCEMENT
	S	S	S
STOCK PRICE STD	0.0200***	0.0135***	
	(7.90)	(4.77)	
STOCK PRICE STD_1			0.00745***
			(3.69)
YEAR DUMMIES	yes	no	yes
COMPANY DUMMIES	yes	yes	yes
CONSTANT	4.299***	15.91***	26.91***
	(19.63)	(280.53)	(139.80)
N	46702	46702	44744

## 4. Key issues in using price target data in analyzing stock options

The goal of my analysis is to use the price target data as a new explanatory variable in analyzing the granting of stock options in executive compensation and to see whether it affects in options in ways consistent with the arms-length bargaining board model of setting executive compensation or with the managerial power model. For price targets to affect decisions about granting executives options, there must be real information in the targets about future share prices. To see if there is information in the target prices I examine their relation to share prices in the forecast period. I also examine whether the stock market responds quickly to announcements as if that investors treat them as informative. For firms to use price targets when boards award stock options, price target announcements and stock option grants must be made in similar periods of time.

The analysis in this section shows that there is information in the price targets that helps predict share prices a year later and that produces changes in market prices shortly after an announcement, and that the timing of announcements has become scheduled in ways that boards can use information in deciding on grants of options.

## 4.1 Is there information in analysts' price targets?

The question of whether or not forecasts of stock prices are useful indicators of future share prices has long been an issue of debate in the financial forecast literature. Using new statistical methods Granger (1992) points out that while there are no general profitable trading rule there is some forecastability in share prices and gives a critical survey of the methods used to forecast price. Most models look at previous firms' performance and use lagged stock prices for forecasting future prices. Such studies indicate that there is a regular pattern in stock prices. Nevertheless, a number of studies point out that shares that performed badly in the first period tend to perform well in the second period. Empirical evidence proves that a price reversal occurs daily, weekly and even monthly. To the extent that price targets are based on "inside information" that analysts pick up from their research, the targets should add forecastability to any models based on common knowledge of share prices.

I examine whether the IBES price targets are predictive of future changes in share prices as follow. I begin with the daily share price reported in CRSP for which IBES provides price targets. For every day analysts announce a price target, I form a triplet of data - the share price on the day the forecast was made, the forecast price a year later, and the actual price a year later (or, when due to timing, that is not available, the share closing price one year and three days later). A triplet for a particular company would be the price target of \$18.00 given on 06/21/2004; the price on that day, say of \$16.95 and the actual share of \$18.26 one year later on 06/21/2005. In 18.65% IBES reports more than one announcement in one day, I thus use the average. I then compute the daily logarithmic return by taking the difference between the logarithm of share stock closing price one year later and the logarithms of the actual stock closing price.<sup>23</sup> I define the expected return the difference between the natural logarithm of the expected share price one year from now and the natural logarithm of the stock closing price.

The null hypotheses for forecastability require that  $\beta \neq 0$  for the fixed effects regression (2.1):

$$R_{t+1,i} = \alpha + \beta E_t (R_{t+1,i}) + \partial yr * + \eta_i + \varepsilon_{t+1,i}$$

$$(2.1)$$

where the dependent variable is the company i daily stock log-return one year later the day t. The independent variable is the expectation at the time the prediction is made for the stock

<sup>&</sup>lt;sup>23</sup> I ignore 0.0049% of the dataset that report an average of zero price targets in order to use the log transformation.

price next year (t+1).  $\eta_i$  is the fixed effect,  $\varepsilon_{t+1,i}$  is the error term, while  $yr^*$  represents the year dummies.

Table 2.7 gives the regression estimates of how analysts' predictions do.

<b>A</b>	(1)	(2)	(3)
	$R_{t+1,i}$	$R_{t+1,i}$	$R_{t+1,i}$
$E_t(R_{t+1,i})$	0.120***	0.181***	0.156***
	(48.17)	(73.78)	(65.21)
YEAR DUMMIES	no	yes	yes
COMPANY DUMMIES	yes	yes	no
CONSTANT	-0.0756***	-0.207***	-0.175***
	(-96.38)	(-61.74)	(-28.04)
N	439250	439250	439250

Table 2.7: Testing analysts' forecast. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%, 5% and 1% levels respectively. The t-statistic are reported in parentheses.

In general, it is reasonable to rely on analysts' predictions since in all specifications  $\beta$  is significantly different from zero and positive. Intuitively, when analysts announce an increase (decrease) in share price, the share price moves in the predicted direction.

The requirement that beta is positive and significantly different from zero is essential for investors to trust analyst's reports. A positive beta indicates that when analysts expect an increase in share price, it actually occurs it. However, more restrictive requirements are needed to prove the unbiasedness of price target. Indeed, the unbiasedness hypothesis requires jointly that the constant is zero and the beta coefficient is equal to one in the ordinary least square regression (Dokko *et al.*, 1989).

In the appendix I test two alternative models to see if it is reasonable for investors to rely on price targets. I find that beta is positive and significantly different from zero in all two models, which implies that it is reasonable for investors expect a rise in the future share price when analysts announce it.

Figure 2.3 illustrates the histogram of the difference between the stock price and the forecast made one year later, which reflects the accuracy of price targets. On average the difference is \$-37.5. The standard deviation is approximately \$9,450 while the median is also more or less \$1.56. However, when I drop 360 observations that forecast a change in the share

price of 10,000% or more, I find that the average is positive (\$2.69), the standard deviation is about \$191.4649 and the median of \$1.55.





An alternative test of the information content in analysts' price target announcements is to see how the stock market responded to the announcements. Womack (1996) studied the market reaction to analysts' recommendations and found that positive recommendations are associated with positive returns in the three days window around the announcement. Asquith *et al.* (2005) also finds that the market reacts in a short period of time to price target announcement.

Establishing the causality between price target and share price is not straight-forward since the factors that cause a positive or negative price target announcement are likely to be the same factors that lead to a rise or a drop of the share price. The key issue, however, is not the underlying cause for the relation but whether the announcement alters market views and serves as the "conduit" for the information. Accordingly, I follow the analyses of Womack (1996) and

Asquith *et al* (2005) and examine how the market reacted to a price target announcements in the IBES data set.

I use a difference in difference approach, regressing the change in share price on the change in the price target for a particular. That is, I take as my dependent variable the stock closing price after a new price target announcement minus the share price after the previous price target announcement and take as my independent variable the difference between the new price target and the earlier price target.<sup>24</sup>

Table 2.8: Testing price targets' effect on share price. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%, 5% and 1% levels respectively. The t-statistic are reported in parentheses. (1)(2)(3) (4)  $sp_{t,i}$  $sp_{t,i}$  $\Delta s p_{t,i}$  $\Delta sp_{t,i}$ 0.786\*\*\* 0.0610\*\*\*  $pt_{t,i}$ (1033.30)(135.31) $sp_{t.i}(-3)$ 0.938\*\*\* (1977.05) 0.0255\*\*\* 0.0254\*\*\*  $E_t(\Delta p t_{t+1\,i})$ (24.54)(24.41) $\Delta sp_{t+1,i}(-15)$ 0.0183\*\*\* (2.97)YEAR DUMMIES yes yes yes yes COMPANY yes yes yes yes **DUMMIES** CONSTANT 0.693\*\*\* 0.00267\*\* 0.00474\*\*\* 0.00473\*\*\* (234.62)(2.85)(2.84)(2.48)536993 Ν 536936 81064 81042

Table 2.8 gives the results of my analysis. In column 1 I regress the log of the share price at the announcement date on the natural logarithms of the price target.<sup>25</sup> Column 2 shows the results of regressing the natural logarithm of share price on the natural logarithms of price target and the natural logarithm of the closing share price three days before a new price target is announced. Column 3 regress the difference of the natural logarithms of the stock closing price at the announcement of the price target with the natural logarithms of the share price target with the log of

<sup>&</sup>lt;sup>24</sup> The results hold also assuming that errors are uncorrelated and adopting the following specifications:  $ln(sp_{t,i}) = \alpha + \beta_1 ln(pt_{t,i}) + \beta_2 ln(sp_{t-1,i}) + \partial yr * + \eta_i + \varepsilon_{t,i} \text{ or } ln({}^{sp_{t,i}}/_{sp_{t-1,i}}) = \alpha + \beta_1 ln({}^{pt_{t,i}}/_{sp_{t-1,i}}) + \partial yr * + \eta_i + \varepsilon_{t,i}$ 

<sup>&</sup>lt;sup>25</sup> I obtain similar results by regressing the log of a stock closing price 15 days and 30 days after the announcement respectively on the natural logarithms of a price target.

the previous price target. (-15) is defined as the logarithms of the stock closing price at the announcement date less the logarithms of share closing price 15 days before the announcement. Intuitively, columns 3 and 4 test if a change in the price target produces a change in the share price. The results show that as soon as analysts release a price target adding new information to a previous price target, the market reacts. The coefficient is positive and significantly different from zero. Therefore, I conclude that a positive price target announcement led to an increase in the share price.

The next question I ask about the price targets is whether or not they are scheduled over time in such a way as to be potentially useful in deciding on granting stock options. I define an announcement as scheduled if the firm has received at least one price target the previous year, and if the actual announcement was done by the same broker within a 7 days window around one-year anniversary of the previous announcement.<sup>26</sup> The first column of table 2.9 shows how many scheduled announcements were made each year. Column 2 gives the number of announcements for which the company has already received at least one price target from the same broker firm last year. Column 3 reports the percentage of announcements for which the company has received at least one price target last year by the same broker company that were also scheduled. As a general trend, the number of scheduled announcements has risen from almost 16% to 30%.

Bradshaw *et al.* (2012) show that target price revision is positively associated with price target accuracy. They define a price target as reviewed if it is issued over one week to six months after a previous price target. Adopting a similar classification and defining a price target as reviewed if the same broker company issued it seven days to 180 days after a previous

<sup>&</sup>lt;sup>26</sup> I treat 242 announcements as scheduled even if no announcements were reported the previous calendar year. For example, an analysts issue no announcements in 1999. In 2000 the analyst issues a price target at the begin of January, and another at the end of December. If the latter one was issued in a window of one week from the announcement made in January, I consider it as scheduled.

announcement, I find that approximately 70% of price targets are a revision of a previous announcement and 20.37% of reviewed price targets are scheduled.<sup>27</sup>

Table	2.9	: Sched	uled prio	ce target's
annou	ncemer	ıts.		
		(1)	(2)	(3)
		#		%
		scheduled	# already	scheduled
	2000	2188	174778	1.25%
	2001	3204	20322	15.77%
	2002	5500	30078	18.29%
	2003	7134	33669	21.19%
	2004	9223	38789	23.78%
	2005	10106	40984	24.66%
	2006	11259	42785	26.32%
	2007	13000	48854	26.61%
	2008	15292	58448	26.16%
	2009	16899	56895	29.70%
	2010	19625	62078	31.61%
	2011	21162	71241	29.70%

However, as shown, 98% of price target announcements have 12-months horizon. Thus, what it is considered a revision actually adds extra information. Indeed, if a price target with one-year horizon is announced in January 2000 and a so called 12-month forecast "revision" is made in March 2000, the new price target actually add three extra months to the previous forecast. The new price target announcement might be considered a revision if the horizon is nine months, since it would refer to the same forecast period previously considered. Because I focus exclusively on 12-months horizon price target revision. However, it is still true that many announcement and a price target revision. However, it is still true that many announcements update previous forecasts. The frequency with which price targets are updated increases the accuracy of a price target, but it might also undermine the credibility of analyst' price targets.

 $<sup>^{27}</sup>$  Reviewed price target were roughly 40.04% in 1999, 57,59% in 2000, 58.25 in 2001, 66.32 in 2002, 68.28% in 2003, 70.63% in 2004, 70.16% in 2005, 71.73% in 2007, 75% in 2008, 77.91 in 2009, 72.3% in 2010 and 74.96% in 2011, while the announcements that were reviewed and scheduled at the same time grew from 12.1% in 1999 to 29.56% in 2011.

# 6. Conclusion

This chapter has introduced the IBES History Price Targets data file from 1999 to 2011 - a data set that has been available for academic research beginning in 2009 - which I use to test the arms-length bargaining and managerial power models of setting the compensation of CEOs. It shows that 98% of price targets have one-year horizon; that the number of announcements depends positively on the stock price volatility; that the percentage of scheduled announcements has risen from 15% in 2001 to almost 30% in 2011; and that price target are generally accurate and thus that investors can rely on them. The evidence that positive (negative) announcements are followed shortly thereafter by an increase (decrease) in share price implies that investors incorporate analysts' information in valuing firms. Do boards and CEOs do the same? The next two chapters show that they do and in ways that cast light on the arms-length bargaining and managerial power models of setting CEO compensation.

### Appendix 2.A: Price targets accuracy: two different models

In this appendix I adopt two different models to test if the price targets are accurate. I show that in all two models investors can rely on price targets.

The first model tests whether a shocks in the price targets is followed by a shock in the share price. In particular, I test if a change in prediction is followed by a change in the share price. More formally, I test the following model:

$$\Delta sp_{t+1,i} = \alpha + \beta E_t (\Delta p t_{t+1,i}) + \partial yr * + \eta_i + \varepsilon_{t+1,i}$$
(2.2)

where  $\Delta sp_{t+1,i}$  is the difference of natural logarithms of the effective stock closing price one year after the first 12-months forecast less the natural logarithms of first effective stock closing price.  $E_t(\Delta pt_{t+1,i})$  is similarly defined as the difference of natural logarithms of the price target less the natural logarithms of the previous price target.

	(1)	(2)
	$\Delta sp_{t+1,i}$	$\Delta sp_{t+1,i}$
$E_t(\Delta pt_{t+1,i})$	0.00222**	0.00213**
	(2.57)	(2.47)
YEAR DUMMIES	yes	no
COMPANY DUMMIES	yes	yes
CONSTANT	-0.00614***	-0.000467***
	(-4.66)	(-2.67)
N	63641	63641

Table 2.10: Testing analysts' forecast. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%,5% and 1% levels respectively. The t-statistic are reported in parentheses.

A second approach directly tests the price level. More specifically,

$$sp_{t+1,i} = \alpha + \beta E_t (pt_{t+1,i}) + sp_{\overline{t},i} + \partial yr * + \eta_i + \varepsilon_{t+1,i}$$

$$(2.3)$$

where  $sp_{t+1,i}$  is the natural logarithm of the stock closing price one year after the forecast is made and  $E_t(pt_{t+1,i})$  is the natural logarithm of the price target announced at time t.  $sp_{\bar{t},i}$  is the natural logarithm 15 (30) days before the announce was done. Indeed, as the literature pointed the price target is positively correlated with the share price. Thus, in order to control for correlation errors I lag the stock closing price, taking the stock closing price 15 (30) days before the announcement.

	(1)	(2)	(3)	(4)	(5)
	$sp_{t+1,i}$	$sp_{t+1,i}$	$sp_{t+1,i}$	$sp_{t+1,i}$	$sp_{t+1,i}$
$E_t(pt_{t+1,i})$	0.375***	-0.0217***	0.0528***	0.102***	0.349***
. ,.	(270.00)	(-9.75)	(24.20)	(48.38)	(244.52)
$sp_{\bar{t},i}(0)$		0.513***			
		(221.53)			
$sp_{\bar{t},i}(15)$			0.433***		
			(187.06)		
$sp_{\bar{t},i}(30)$				0.381***	
				(168.77)	
YEAR	yes	yes	yes	yes	20
DUMMIES					IIO
COMPANY	Ves	yes	yes	yes	yes
DUMMIES	yes				
CONSTANT	1.942***	1.521***	1.542***	1.556***	1.972***
	(346.41)	(269.50)	(265.73)	(263.51)	(420.94)
Ν	439375	439250	439050	437682	439375

Table 2.11: Testing analysts' forecast. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%, 5% and 1% levels respectively. The t-statistic are reported in parentheses.

Table 2.11 reports the results. Column two shows that the price target coefficient switch from positive to negative, if I introduce the stock closing price at the announcement day. This is due to the positive correlation between the share price and the price target at the announcement, which let the errors be correlated. One way to control it is to lag the stock closing price. In column three I lag the stock closing price of 15 days while in column four I lag it of 30 days. Once controlling for the errors correlation, the results show that price targets are positively associated with effective stock price at the forecasted date and thus investors can rely on analysts' forecasts.

## Chapter Three: The Effect of Price Targets on the Composition of CEO Pay

## 1. Introduction

In this chapter I analyze the impact of an indicator of the likely future change in a firms' stock price on executive compensation. The indicator is the expectations that market analysts give for the firms' share price in the next year as reported by the Institutional Broker Estimate System in its Detail Price History Target data base, as described in chapter two. The share price that analysts expect for a firm in the future has not to my knowledge been previously analyzed in studies of executive compensation. I show that its relation to the composition of executive compensation provides insights into the determinants of executive pay.

My main finding is that analysts' forecasts of the share price affect the structure of CEO compensation. I then examine how well the two main models of the determination of executive pay - the arm's length bargaining model, which assumes the board of directors contract top executives pay to maximize shareholders' interests; and the managerial power model, which assumes the board of directors maximizes managers' compensation within constraints imposed by social costs and market penalties (Bebchuk *et al*, 2005), account for estimated relations between expected share prices and compensation. Reviewing empirical studies, Frydman *et al*. (2010) conclude "both managerial-power and competitive market forces are important determinants of CEO pay" but leave open the question of their relative importance. Extant analyses of the substitution between cash-based and stock-based pay in executive compensation have not analyzed the effect of analysts' expected changes in share prices on that substitution nor sought insight into the arm's-length bargaining and the managerial power models from such data.

As described in chapter two, the IBES Detail History Price Target data file reports

companies' price targets from different analysts. Analysts' announcements are public information available to the board of directors and to shareholders as opposed to private information that managers can use to structure compensation for their benefit at the expense of shareholders. Managers may influence analysts' recommendations and/or anticipate them but public information gives them less room to manipulate their compensation than private information.

I find that analysts' price targets alter the composition of executive pay between cashbased compensation and stock-based compensation. When analysts forecast a rise in the share price for a firm, its compensation package tilts toward stock-based compensation. When analysts forecast a fall in the share price, the compensation package tilts toward cash-based compensation. This pattern is more readily explicable by the managerial power model than the arm's length bargaining model. Consistent with this interpretation, the trade-off is stronger in companies that have weaker corporate governance.

# 2. Background

Shareholders hire executives to act in their interest. But managers have information that shareholders do not have, and they can exploit that information for their benefit. Thus, managers' interests are generally not fully aligned with shareholders' interests. The principal-agent theory suggests that the agency cost rising from separation of ownership and control can be mitigated through stock-based compensation that links executive pay to the share price of the firm, which reflects company performance.

The ideal board model of setting executive compensation posits that the firm's board of directors bargain at arm's length with executives for a stock-based incentive scheme that ideally induces management to maximize shareholders wealth. Motivated by the desire to better align

compensation and shareholder value, major firms shifted executive pay from cash-based salaries to stock-based compensation, primarily stock options, from the 1980s to the 2000s. In the late 2000s, they began to shift stock-based compensation from options to restricted stock grants.

The managerial power model questions the efficacy of stock-based incentives to solve the principal agent models and views incentive schemes as part of the agency problem itself. It argues that boards of directors whose members may be in part chosen by executives and in which executives serve do not bargain ideally on behalf of shareholders. Bebchuk *et al* (2003) explore how the two models influence top executive compensation and direct attention at aspects of executive pay determination that the arm's-length bargaining cannot explain. They argue that the observed pay structure is most likely a compromise between market forces, which mitigates managerial rent-seeking, and managerial-power that favors top executive compensation. Analyzing the growth of manager compensation during the period 1993-2003, Bebchuk *et al.*, (2005) note that equity-based compensation, consisting largely of stock options, increased considerably without any reduction of cash-based compensation. This raises the possibility that the complexity of options may have given executives a way to increase their compensation.

There are measurement and modeling issues in comparing the value of stock options to salary or other cash-based pay. Hall *et al.* (2002) have argued that the Black-Scholes formula traditionally used to value stock option compensation overstates the value of options granted to executives. The Black-Scholes method assume that the holders of options are risk neutral inasmuch as they can hedge the risk by short selling and diversifying their portfolio. If managers cannot hedge the risk because they have to keep their compensation aligned with firms' performance, they face a higher risk-related cost for exchanging cash-compensation for options. Murphy (2002) claims that executives are willing to exchange cash for options but for a risk premium.

On the other side, Bettis *et al.* (1999) give evidence that executives hedge some options. There are, moreover, other forms of compensation policy that mitigate the opportunity cost of holding stock instead of cash. When options go underwater companies often seek to re-align incentives by granting extra options (makeup grants), or by deleting underwater options and granting new ones in a "6&1 exchange" no earlier than six months and one day, where the timing reflects accounting rules. Less common is the practice of repricing options by reducing the exercise price of existing underwater options. There is a danger that such re-alignment policies may have adverse effects on incentives since they compensate executives despite poor performances. To the extent that firms adopt such policies, they reduce the risk of receiving stock-based compensation instead of cash, which mitigates against the Hall et al critique.<sup>28</sup> Balachandran *et al.* (2004) find that firms are more likely to grant extra options without changing salaries when the firm performs poorly and existing options go under water.

The shift from cash-based compensation to options is also impacted by the tax advantages to the firm of paying large amounts as stock options rather than as cash.<sup>29</sup> Under Section 162m of the US tax code a firm can deduct incentive pay to executives over \$1 million as a cost of business but it cannot deduct a salary over \$1 million as a cost of business. Since paying executives in a way that minimizes taxes is in the interest of shareholders, an optimal board may choose to pay with options even if the options do not truly align management and shareholder interests. As some executives have gained massively from options even when their firm is doing poorly, shareholders and the public have become more critical of options.

<sup>&</sup>lt;sup>28</sup> As Zamora (2005) points out some scholars argue that firms adopt such responses as alternative source of financing. Indeed, by replacing cash-based compensation by stock-based, firms may finance other project in the short run. However, Zamora (2005) find evidence that companies do not recur to options grants to alleviate cash flow problem.

<sup>&</sup>lt;sup>29</sup> There are two kinds of options: Non-Qualified stock options and Incentive stock options. The first one, allows firm to deduce them from taxes. However, the employee is taxed on the spread between market and exercise price once he or she exercise them. The advantage for the employee is that there is no holding requirement. On the other hand, Incentive Stock Options confer a better tax treatment to executives but impose the holder to hold onto the stock for a longer period, which raises the risk. Plus, firm cannot deduce the gain as compensation expense. The not holding requirement and the deducibility explains the popularity of Non-Qualified options.

Kuhnen *et al.* (2012) shows that in the recent shift out of stock options to less contentious form of compensation such as salary is associated with increased press negativity toward options.

In short, there are several factors associated with the allocation of compensation between cash-based and stock-based pay: shareholder desire to align management interests with their own; tax advantages of compensation beyond a million dollars, publicity and shareholder concern over excessive pay, and the risk trade-off to executive of compensation in options to compensation in cash.

#### 3. Data

My analysis uses the two data sets on executive compensation described in chapter one and the IBES price target data described in chapter 2 to examine how analysts targets affect the composition of compensation.

The first data set on executive compensation is Standard & Poor's Executive Compensation data set (Execucomp) for fiscal year 1999 through fiscal year 2010, as described in detail in chapter one of this thesis. I focus my analysis on CEOs. I take the following information from Execucomp: the dollar value of salary (Execucomp variable SALARY in thousands of dollars), bonus (Execucomp variable BONUS in thousands of dollars), the value of stock options awarded (Execucomp variable BLK VALUE until 2006, rename as AWARD\_FV thereafter), the value of stock awards (Execucomp variable RSTKGRANT before 2006, renamed STOCK AWARD FV thereafter), the Long Term Incentive Payout (LTIP), Deferred Compensation Earnings Reported as Compensation (DEFER\_RPT\_AS\_COMP\_TOT, available only after 2006) and All Other Total (ALLOTHTOT substituted in 2006 by OTHCOMP). I also collect the number of derivatives granted to each manager (Execucomp variable OPTION\_AWARD\_NUM). In 20 cases

Execucomp indicates two officers as CEOs for the same firms, which become 9 in the final data set (once I drop firms for which was not possible assign a PERMNO). I therefore use the mean.

To obtain detailed information on stock options, I use data from the Thomson-Reuters Insider Filing Data Feed (IFDF) Table II, as described in chapter one. I obtain the number of options granted, the official grant date, the exercise date, the expiring date and the exercise price of options awarded to company's CEO, which allow inferring the percentage of options scheduled, unscheduled or backdated. I merge the IFDF and the fiscal year Execucomp data using the PERMNO code for a firm. The IFDF provides me with information that I use to estimate the percentage of scheduled, unscheduled and backdated options. I drop firms for which I could not match the data sets by the PERMNO code.

Finally, I link this compensation data to the price target data obtained from the Institutional Broker Estimate System (IBES) Detail History Price Target data file. The IBES data records analysts' price targets level, analysts' name, the company he/she works for, company for which he/she issues the target price, the horizon-period, the day the price target was announced and when it became active in IBES data file, the company currency and whether or not the company is a US firm. I limit my analysis to US firms.

# 4. Trade-off between cash-based compensation and stock-based compensation

The first step in my analysis is to examine whether the board of directors takes into account price targets when setting CEO pay. The null hypothesis is that the board sets CEO compensation independently from the likely course of company share price. There are two alternative scenarios. The first is that the firm shifts CEO pay from cash-based compensation to stock-based compensation when prices are expected to rise, which makes options more desirable and shifts CEO pay from stock-based compensation to cash-based compensation when prices are expected to fall, which makes options given at the money less desirable to the CEO. This strategy benefits executives. The alternative policy would be for the firm to shift CEO pay away from stock-based compensation when it expects the share price to rise. Since the options are expected to be more valuable in the future, a smaller number would give the executive the same value and incentive.

I focus on IFDF companies for which IBES issue price targets and Execucomp reports compensation details.<sup>30</sup> Equation (3.1) defines the empirical model I estimate:

$$DEP_VAR_{i,t} = \propto +\beta_1 EXPROFIT_{t,i} + \beta_2 X_{i,t} + \delta yr * +\lambda_i + \varepsilon_{i,t}$$
(3.1)

 $DEP_VAR_{i,t}$  stands for dependent variable. In this section my dependent variable is the share of the executive's pay package that takes the form of stock-based compensation, SBC, defined as the sum of stock options and stock award divided by total compensation as reported by Execucomp (TDC1). By definition, it is one minus the share of executive pay that is cashbased.

The explanatory variable of interest in this analysis is the fiscal year average of the difference between the price target and the stock closing price, which I label EXPROFIT. I calculate this variable by subtracting from the price target the stock closing price at the announcement day obtained from CRSP (intrinsic value, using the stock options terminology) for IFDF companies for which IBES reports price targets in a fiscal year. I then compute the average difference between the price target and the stock closing price for the fiscal year. I use the Compustat variable FYR, which measures the months on which a given firm's fiscal year ends, from Execucomp, to allocate each closing price and price target to the right fiscal years.

<sup>&</sup>lt;sup>30</sup> In the appendix I discuss the substitution effect for the entire companies reported by Execucomp datasets for which IBES provides price targets.

from 1999 to 2010. If no FYR information is available I assume that fiscal year coincides with calendar year, since <sup>3</sup>/<sub>4</sub> of S&P500 companies have a calendar year fiscal year (Wharton Research Data Services, n.d.). The reason I transform the data onto a fiscal year basis is that the information in Execucomp is on a fiscal year basis.

I also include a vector of measures of the financial and economic characteristics of the firm:

- RETURN: is the rate of return.
- SALE: is the natural logarithm of the gross sales or of the amount of billing for regular sales in thousand of dollars.
- CHSR: is the number of shareholders of ordinary shares/common capital (in thousands dollars).
- EFFICIENCY: is the difference between Tobins Q at time t and Tobins Q at time (t-1). When EFFICIENCY is low, there should pressure on management to restore it.
- FEMALE: a dummy variable taking the value of one if the CEO is a woman. If more than two CEOs are reported in a firms it stands for the percentage of female as CEO.
- AVGAGE: indicate the age of the CEO. If more than one CEO is reported in a company, the item stands for the average age of the officers.
- AT: is the natural logarithm of the current assets in million.
- $yr^*$  stands for time dummies for each fiscal year from 1999 to 2010.
- $\lambda_i$  is the firm fixed effect.

Table 3.1 summarizes the results of an OLS regression of equation 3.1 in which I include company fixed effects. The estimated coefficient on the EXPROFIT variable measures the effect of share price expectancy on CEO stock-based compensation.

Significance at 10 %, 5 % and 1 % reversitespectively. The elsuadate are reported in parenticies.					
	(1)	(2)	(3)	(4)	(5)
	SBC	SBC	SBC	SBC	SBC
EXPROFIT	0.000473***	0.000407**	0.000447***	0.000454**	0.000433**
	(2.85)	(2.44)	(2.71)	(2.10)	(2.01)
SBC_1			0.0339***		0.0365***
			(3.53)		(3.71)
EXPROFIT_1				-0.00000155	-0.00000488
				(-0.01)	(-0.02)
RETURN	-0.0388***	-0.0386***	-0.0386***	-0.0351***	-0.0351***
	(-7.15)	(-7.11)	(-7.11)	(-6.00)	(-6.00)
SALE	0.00321	-0.0277**	0.00270	0.000705	0.000549
	(0.41)	(-2.54)	(0.34)	(0.09)	(0.07)
CSHR	0.00000440	-0.00000870	0.00000944	0.0000166	0.0000215
	(0.08)	(-0.16)	(0.17)	(0.29)	(0.38)
EFFICIENCY	0.00241*	0.00318**	0.00195	-0.000107	-0.000239
	(1.69)	(2.22)	(1.37)	(-0.06)	(-0.13)
FEMALE	0.0349	0.0386	0.0330	0.0335	0.0329
	(1.44)	(1.59)	(1.37)	(1.34)	(1.32)
AVGAGE	0.000801***	0.000804***	0.000795***	0.000721***	0.000694***
	(3.70)	(3.71)	(3.66)	(3.22)	(3.10)
AT		0.0450***			
		(4.09)			
YEAR	yes	yes	yes	yes	yes
DUMMIES					
COMPANY	yes	yes	yes	yes	yes
DUMMY					
CONSTANT	0.395***	0.286***	0.383***	0.418***	0.403***
	(6.78)	(4.49)	(6.56)	(6.91)	(6.65)
N	11909	11909	11823	11394	11317

Table 3.1: Multivariate analysis of substitution between cash-based compensation and stock-based compensation for CEOs for the period from 1999 to 2010. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%, 5% and 1% levels respectively. The t-statistic are reported in parentheses.

The estimates show first that price target matters in compensation. In all of the calculations, the expected future share price enters with a substantial and significant positive coefficient. There is a strong trade-off between cash-based compensation (salary, bonus and "others compensation") and more risky stock-based forms of reward (stock options and stock award). The higher the expected profit (measured as the difference between the future price and the stock closing price) the higher is the share of CEO pay given in stock or stock options. Some of other regression coefficients also deserve attention. Older CEOs receive a greater share of pay in stock-based compensation. This might due to the fact that they have more assets than younger CEOs and thus are less risk-adverse. In the regression that includes measure of total asset, the variable for sales is significantly positive. This might due to the fact that both

sales and assets reflect the size of the firm or that CEOs view a high sales to asset ratio as an indicator of future share prices.

Could the relation between the analysts' target price relative to the stock closing price at the time the target price was given and the stock-based share of compensation result from a reverse causality? Stock options are granted with the purpose of giving the CEO incentives to undertake policies that raise share price. It is possible that analysts take into account CEOs' compensation when issuing price targets on the assumption that a greater stock-based share of compensation does in fact raise the share price a year later. But such an interpretation runs into two difficulties. First, there is no compelling evidence that greater stock options are associated with future improvements in share prices. The empirical work shows that options produce a strong positive contemporaneous relation between share prices and executive compensation, not that greater options in t are associated with a future increase in share prices from t to t+1. Second, the analysts would have to know the compensation plans of the firm before they made their price target, which is highly dubious, as compensation committees are unlikely to divulge to outsiders their intended pay to their CEO. Third, even if the analysts knew company plans for regular scheduled options or extrapolated previous option grants, they cannot plausibly know about company plans unscheduled options that are, by definition, granted at irregular interval. In the appendix, I show that price targets are positively associated with unscheduled options but not with scheduled ones, and thus that I can exclude a reverse causality between price targets and stock-based compensation.

### 4.1 Price targets and Non-equity incentives and LTIP

A component of CEO's cash-based compensation that has received little analytic attention in the executive compensation literature is payments labeled "non-equity incentives".

Non-equity incentives differ from other forms of cash-based compensation because executives have to meet certain pre-established and disclosed criteria to obtain them. The SEC introduced the term "non-equity incentives" to define cash incentive awarded if executives meet a disclosed pre-established target. LTIP are earned if CEO satisfies criteria measured over a time period longer than one year, typically three to six years (Larcker, 1983). The firm usually expresses the target in terms of earnings per shares set at the beginning of the award period. Execucomp reports LTIP until 2006 when it constituted approximately 7% of CEOs' total compensation (TDC1). Execucomp did not report when performance goals were established and thus when price targets could have affected LTIP compensation. Before 2006 Execucomp defined cash-based compensation beyond salaries as a bonus but in 2006 it added the nonequity incentive category and limited bonuses to cash compensation given for reaching undisclosed goals. The result is a discontinuity in the Execucomp data. In 2005 bonuses were 21% of total compensation whereas in 2006 bonuses were less than 10% of total compensation. But LTIP, bonus, and non-equity incentives constitute on average 27% of CEO's total compensation, which suggest that the categories as a group basically reflect the old bonus category.

The value earned as LTIP and non-equity incentive is disclosed when the executive satisfies the performance criteria, not when it sets the criterion. This is very different from stock options, where the firm discloses the options when granted. Price targets might affect non-equity incentives when the firm sets the criteria for gaining the extra pay for meeting the criteria. But price targets cannot affect non-equity incentives when the firm reports the outcomes. To the extent that price targets affect the compensation labeled as non-equity incentive, it must do so with a lag. Balachandran *et al.* (2010) report that in most cases, firms set non-equity incentives annually based upon a single year targets. Therefore, to estimate the possible impact of price targets on non-equity incentive, I examine the relation between non-equity incentives and EXPROFIT lagged one year.

Even with the lag, there could be a positive simultaneous correlation (but not causality) between the expected share price and non-equity incentive. For example, suppose that a company sets a non-equity incentive at time t-s. If at time t the sector is booming, the executive will likely meet the performance criteria. If analysts issue a high price target for the company because they expected the boom to continue, we would obtain a positive correlation between share price and non-equity incentives even though there is no causal relation between them. A future price target cannot affect the terms of a non-equity incentive that were set a year earlier.

To estimate the impact of price targets on the new variable "non-equity incentive" I proceeded as follow. First, I defined the dependent variable as the value of non-equity incentive divided by total compensation. I then lagged the main independent variable EXPROFIT by one year on the notion that the firm could have taken account of that information in setting the terms that would determine non-equity pay a year later.

The results in table 3.2 show a positive insignificant relation between price targets and non-equity incentive in the same period. When I add a measure of the rate of return to reflect the company's performance, arguably a better indicator of the likelihood that the executive would have met the criterion for the non-equity incentive, the magnitude of the estimated coefficient on EXPROFIT decreases. But the key estimated parameter in the table is the coefficient linking the lag of EXPROFIT and non-equity incentives. It is significantly negative. This confirms the table 3.1 finding that firms shift away from cash-based compensation for CEOs when the share price is expected to rise, even from non-equity incentives. They trade non-equity incentive for forms of compensation, like stock options, that are more highly leveraged to increases in share prices.

Why does the firm treat non-equity incentives different from stock-based incentives? One might argue that CEOs regard non-equity incentives as stock options since they are earned only if a pre-set criterion is met. The higher the price target is, the higher the probability of reaching the goal so that the CEO would gain from both non-equity incentives and stock options. However, there is a major difference between options and non-equity incentives. The intrinsic value of an option depends exclusively on the share price: if the share price is above the strike price the option is in the money. If the price is below the strike price, the option run out of the money and cannot be cashed. Non-equity incentives, by contrast, are not exclusively linked to share price and may indeed relate to other performance goals. If the share price is expected to fall, a CEO might convince the board to set the criteria for her/his non-equity incentives based upon goals other than the share price - for instance market share or increases in sales. Finally, non-equity incentives are tax-attractive for companies. Firms can qualify for the deduction under the section 162 of tax code as long as they meet a pre-established performance goal of almost any kind, including a "performance goal need not, however, be based upon an increase or positive results under a business criterion" (Section 162(m) of the Internal Revenue Code). If a firm's share price is expected to fall, the firm would set a criteria for a non-equity incentive to be a smaller drop in the share price than analysts expected.

The results in table 3.2 show a positive insignificant relation between price targets and non-equity incentive (NON- EQUITY INC) in the same period. When I add a measure of the rate of return to reflect the company's performance, arguably a better indicator of the likelihood that the executive would have met the criterion for the non-equity incentive, the magnitude of the estimated coefficient on EXPROFIT decreases. But the key estimated parameter in the table is the coefficient linking the lag of EXPROFIT (EXPROFIT\_1) and non-equity incentives. It is significantly negative. This confirms the table 3.1 finding that firms shift away from cash-based compensation for CEOs when the share price is expected to rise, even from non-equity incentives. They trade non-equity incentive for forms of compensation, like stock options, that are more highly leveraged to increases in share prices.

Why does the firm treat non-equity incentives different from stock-based incentives? One might argue that CEOs regard non-equity incentives as stock options since they are earned only if a pre-set criteria is met.

		e reperieu in p				
	(1)	(2)	(3)	(4)	(5)	(6)
	NON-	NON-	NON-	NON-	NON-	NON-
	EQUITY	EQUITY	EQUITY	EQUITY	EQUITY	EQUITY
	INC	INC	INC	INC	INC	INC
EXPROFIT	-0.0000993				0.000385	0.000140
	(-0.37)				(1.04)	(0.32)
EXPROFIT_1		-0.0008***	-0.0013***	-0.0014***	-0.0015***	-0.0016***
		(-3.02)	(-4.46)	(-4.57)	(-4.53)	(-4.36)
NON-						
EQUITY						-0.0967***
INC_1						
						(-6.04)
RETURN			0.0672***	0.0671***	0.0700***	0.0656***
			(13.22)	(13.22)	(13.46)	(12.62)
SALE			0.0577***	0.0546***	0.0575***	0.0537***
			(5.19)	(5.95)	(5.12)	(4.30)
CSHR			-0.000176*	-0.000178*	-0.00018**	-0.0000870
			(-1.91)	(-1.93)	(-1.98)	(-0.75)
EFFICIENCY			-0.00658**	-0.00643**	-0.00700**	-0.0098***
			(-2.06)	(-2.02)	(-2.15)	(-2.97)
FEMALE			0.0518**	0.0517**	0.0526**	0.0592*
			(1.97)	(1.97)	(1.97)	(1.83)
AVGAGE			0.000329	0.000329	0.000342	0.000629
			(1.13)	(1.13)	(1.16)	(1.05)
AT			-0.00602		-0.00775	0.00590
			(-0.49)		(-0.62)	(0.40)
YEAR	yes	yes	yes	yes	yes	yes
DUMMIES	•	2	•	•	•	•
COMPANY	yes	yes	yes	yes	yes	yes
DUMMY	-			-		-
CONSTANT	0.182***	0.220***	-0.209**	-0.233***	-0.197**	-0.275**
	(47.80)	(56.01)	(-2.42)	(-3.24)	(-2.22)	(-2.48)
Ν	7236	6804	5889	5889	5821	4736

Table 3.2: Multivariate analysis of the effects of price targets on non-equity incentive for the period from 2006 to 2010. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%, 5% and 1% levels respectively. The t-statistic are reported in parentheses.

The higher the price target is, the higher the probability of reaching the goal so that the CEO would gain from both non-equity incentives and stock options. However, there is a major difference between options and non-equity incentives. The intrinsic value of an option depends exclusively on the share price: if the share price is above the strike price the option is in the money, but if the price is below, the option run out of the money and cannot be cashed. Non-equity incentives, by contrast, are not exclusively linked to share price and may indeed relate to other performance goals. If the share price is expected to fall, a CEO might convince the board to set the criteria for her/his non-equity incentives based upon goals other than the share price - for instance market share or increases in sales. Finally, non-equity incentives are tax-attractive for companies. Firms can qualify for the deduction under the section 162 of tax code as long as
they meet a pre-established performance goal of almost any kind. If a firm's share price is expected to fall, the firm would set a criteria for a non-equity incentive to be a smaller drop in the share price than analysts expected.

#### 5. Corporate governance structure

The analysis thus far has treated all firms the same, as if the two forces at work in determining executive compensation - arms-length bargaining by boards vs. managerial rent-seeking - had the same influence on pay in all firms. Using data on the structure of corporate governance, I examine next whether firms with stronger or weaker governance evince larger substitution between cash-compensation and stock-compensation in response to analyst forecasts of share prices. Companies that act according to the arm's length bargaining model are less likely to substitute stock compensation for cash compensation when analysts expect share prices to fall than companies that act according to managerial power model. CEOs facing poor corporate governance have more room for rent-seeking than CEOs working for companies with stronger corporate governance.

As my measure of corporate governance I used data on whether a CEO is or is not a member of the board of directors. Thus, the CEO who is also member of the board has a conflictual position inasmuch as she/he has to evaluate her/his performance and can influence the compensation committee on her/his pay. I create a variable (BOARD) that indicates if the CEO has served as member of the board of directors during the fiscal year. Execucomp reports a true/false variable that indicates whether or not the CEO has served as a director. However, this is not enough to define a CEO as a member of the board. I define a CEO as member of the board if the CEO title contains the strings: "chmn", "Chmn", "chairman" or "Chairman". In 228 cases Execucomp reports that CEOs have the title of the board but do not indicate whether

he/she served as director. I consider someone with those titles as being a member of the board of directors. In 7,660 cases the CEO served as director but was not indicated as Chairman. I do not consider them as members of directors.<sup>31</sup> If more than one CEO is listed in the same company for a given fiscal year, I take the sum of CEOs that are members of the board. Almost 78% of the firms report that CEO has been appointed member of the board for at least one fiscal year. However, the percentage of firms with CEOs on the board of directors has progressively diminished from 78% in 1996 to roughly 58% in 2010, which still constitutes the majority.

Using BOARD as an indicator of poor corporate governance, I divide the dataset into companies that have a CEO member of the board and companies that do not. I then run the same regressions as in tables 3.1 and 3.2. Table 3.3.1 shows the coefficient and t-statistic on the EXPROFIT variable using specifications (1) to (4) as in table 3.1 for the two groups of companies. The estimates show that indeed corporate governance matters. Firms having its CEO as a member of the board substitute cash-based compensation with stock-based compensation according to price targets. By contrast, the minority of firms that have stronger corporate governance do not set CEO compensation according to price targets. Similarly, table 3.3.2 reports the coefficient of one period lag of EXPROFIT (EXPROFIT\_1) for specification (2) to (5) in table 3.2 for both subgroups. The results shows that in both subgroups CEOs shift away from cash-based compensation when the price target is expected to rise, but the phenomenon is stronger in companies with weaker corporate governance.

<sup>&</sup>lt;sup>31</sup> Execucomp does not offer variables indicating whether a CEO is or not a member of the board (Wharton Research Data Service, n.d.). According to, WRDS the best way to do it is to look whether the CEO title contains the word "chairman".

Table 3.3: Substitution between cash-based compensation and stock-based compensation for CEOs for the period from 1999 to 2010 by CEO subgroups. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%, 5% and 1% levels respectively. The t-statistic are reported in parentheses.

Table 3.3.1		SBC		
	specification	EXPROFIT	t- statistics	Ν
CEO non member of the board	(1)	0.000264	(1.26)	3820
	(2)	0.000201	(0.96)	3820
	(3)	0.000285	(1.37)	3787
	(4)	0.000174	(0.63)	3648
	(5)	0.000193	(0.70)	3619
CEO member of the board	(1)	0.000674**	(2.22)	8088
	(2)	0.000635**	(2.09)	8088
	(3)	0.000633**	(2.09)	8035
	(4)	0.000911**	(2.35)	7745
	(5)	0.000842**	(2.17)	7697

Table 3.3.2	Non-equity incentives				
	specification	EXPROFIT_1	t- statistics	Ν	
CEO non member of the board	(2)	-0.000574	(-1.25)	2614	
	(3)	-0.000908*	(-1.67)	2259	
	(4)	-0.000912*	(-1.70)	2259	
	(5)	-0.000827	(-1.42)	2225	
	(6)	-0.00106	(-1.60)	1864	
CEO member of the board	(2)	-0.000909***	(-2.85)	4189	
	(3)	-0.00155***	(-4.19)	3629	
	(4)	-0.00162***	(-4.41)	3629	
	(5)	-0.00177***	(-4.42)	3595	
	(6)	-0.00184***	(-3.97)	2872	

#### 8. Discussion

The findings that firms switch from cash-based compensation to stock-based compensation of CEOs when analysts expect their share price to rise, and switch from stock-based compensation to cash-based compensation when analysts expect share prices to fall, and the stronger pattern in companies with weaker corporate governance can be interpreted under both the managerial power model and the arm's length bargaining model. Under the managerial

power model, a CEO will always seek ways to increase his or her income. They will try to exchange cash for stock-based compensation when they expect the share price to rise, whatever the causes of the rise, and prefer cash compensation when they expect the share price to fall. Hence, a board dominated by the CEO and management would always substitute cash for stock-based compensation according to share price targets.

Whether a board of directors acting under the arm's length bargaining model would trade-off cash-based compensation and stock-based compensation according to price targets is less clear-cut. Suppose that the board of directors is fully able to screen CEO's skills. One might argue that, being able to determine CEO's value, the ideal board would simply seek the most efficient way to reward the CEO either using cash or stock compensation accordingly to the contingencies, such as the tax law. Switching from cash-based compensation to stock-based compensation when the share price is expected to rise, the board would save shareholders from paying some costs. For instance, the board of directors may "inflate" the options' value and made it up to a desired amount by using price target information and take advantage of section 162m of tax code for gaining a favorable tax treatment.

But the ideal board would still want to incentivize CEOs by bonding CEO compensation to share price. The principal-agent arguments still remain: no matters how much manager's skills are worth, the board needs to align shareholders' interests to CEO's personal interests. Therefore, the ideal board still needs to determine the right stock-based compensation that maximizes CEO incentives and promote shareholders' interests independently from the cash-based compensation. Under the arm's length bargaining model, stock-based compensation is efficiently set to maximize CEOs' incentives independently from cash-based compensation. From this perspective an ideal board should not contemplate a trade-off between stock-based compensation.

Those companies lacking strong corporate governance are more likely to shift from cash-based compensation to stock-based compensation in response to analysts announcing a positive price target for their shares favors the managerial power model explanation of the substitution effect between cash and stock according to price target. Indeed, the managerial power model works under the hypothesis that managers can influence the board at their own advantage. Obviously, a CEO that is also member of the board has more room to guide board's decision at her/his own advantage and take advantage of a rise in share price by substituting cash with stock-based compensation. That the minority of firms that do not admit its CEO as part of the board set their CEO pay independently from price targets nor shift from cash to stock when it is most suitable for the CEO implies that the observed behavior reflects the influence of CEOs on boards. I thus conclude that the substitution effect between cash-based compensation according to the price targets is a strong argument in favor of the managerial power model as a model that mostly explain CEO compensation. In the appendix I give also evidence that CEOs are granted extra–unscheduled options when analysts expect a rise in share price.

In sum, the relation between IBES reports on analysts' price target announcements and the stock-based share of executive compensation - increases in the stock-based share of compensation when analysts expect a rise in share price, and decreases in the stock-based share of compensation when analysts foresee a drop in company share price are more aligned with the managerial power model of executive pay than with the arm's bargaining model.

### Appendix 3.A: Scheduled vs. unscheduled options

In this appendix I decompose the options granted into those granted at approximately the same time as the previous that I have called scheduled options and those given at some other time, which I have called unscheduled options. I estimate the two types of options react to price targets announcements. Because scheduled stock options are granted with a regular pattern, while unscheduled options are given at irregular times, it is reasonable to expect a different relation to price targets announcement. A CEO may seek more options when a positive price target announcement is released but when there are no options scheduled in the period, the CEO will have to press the board for unscheduled options. This is likely to be associated with a management dominated board. An arms-length bargaining board, by contrast, might bargain with the CEO when setting scheduled options, taking account of recent price target announcements. The board might seek to give fewer options to save the shareholders money since each option would have greater value given the announcement while the CEO would seek for the same reason. As an increasing number of options are scheduled, the key issue would be the number of derivatives included in the scheduled options package when analysts expect a rise in share price.<sup>32</sup>

Table 3.4 reports the distribution of options granted around price target announcements. To construct the table, I looked to see if any price target announcements were reported seven days before or seven days after any options transaction in IFDF. If more than one announcement falls into the 14-days window, I retain only the announcement closest to the options awarding day. I then take the difference between price target and stock closing price at the announcement day.

<sup>&</sup>lt;sup>32</sup> One possible explanation I also explored is whether the board schedules options granting around scheduled price target announcements. However, I did not find consistent results.

14010 0111	Distribution	or options at out a p	fice targets a	# Options
	Negative	Independently	Positive	Granted (in Mio)
	Regative	Total Option		14110)
1999	4 45%	89 19%	6 36%	385 67
2000	12.19%	71.64%	16.17%	518.32
2001	6 49%	72.97%	20 54%	628.16
2002	7.44%	70.39%	22.17%	553.08
2003	9.17%	69.15%	21.68%	546.18
2004	9.92%	68.94%	21.14%	518.89
2005	8.99%	69.30%	21.72%	489.51
2006	6.47%	61.73%	31.80%	455.88
2007	11.12%	60.93%	27.95%	429.27
2008	5.33%	60.11%	34.56%	511.37
2009	5.78%	45.92%	48.30%	659.23
2010	5.87%	59.15%	34.99%	364.67
		Scheduled Opt	ions	
1999	4.83%	89.66%	5.51%	70.42
2000	26.20%	56.01%	17.79%	94.64
2001	10.87%	64.51%	24.62%	113.15
2002	12.95%	59.69%	27.36%	116.89
2003	12.03%	55.55%	32.43%	113.48
2004	20.18%	52.42%	27.40%	113.14
2005	17.37%	48.82%	33.81%	115.58
2006	11.36%	49.63%	39.02%	100.59
2007	10.84%	46.38%	42.78%	100.82
2008	10.45%	39.21%	50.34%	144.7
2009	11.77%	38.24%	49.99%	172.95
2010	7.74%	42.13%	50.14%	139.82
		Unscheduled Op	otions	
1999	7.23%	82.18%	10.59%	152.19
2000	8.28%	74.14%	17.58%	231.45
2001	7.12%	68.58%	24.30%	252.74
2002	6.36%	68.20%	25.44%	257.99
2003	9.43%	70.23%	20.34%	263.38
2004	7.30%	71.55%	21.15%	237.46
2005	7.08%	72.76%	20.16%	213.55
2006	5.32%	51.49%	43.19%	195.8
2007	8.29%	63.98%	27.73%	169.88
2008	5.00%	64.85%	30.15%	194.62
2009	5.35%	63.31%	31.35%	215.06
2010	6.54%	59.93%	33.54%	113.26

Table 3 1: Distribution of options around price targets appearance

Finally, I sum the total number of options granted around a positive, or negative price target announcement as well as those options granted independently from price targets by fiscal year.

Table 3.4 shows that the percentage of options granted around positive announcements has grown substantially, which emphasizes once again how price targets play a growing role in setting CEO compensation. The percentage of both unscheduled and unscheduled options granted around positive announcements has also grown even though unscheduled options have generally decreased since 2001. Only a small percentage of options are granted around negative price targets announcement.

To study the relation between price targets and scheduled and unscheduled options, taken separately, I merged IFDF data on the total number of unscheduled options granted in a fiscal year and on scheduled options with data from Execucomp by fiscal year and companies. Because Execucomp has so little information on options granted, it is not possible to inferring from it whether options are scheduled or unscheduled. For consistency with the table 3.1 calculations, I define the amount of scheduled options as the percentage of scheduled option inferred from IFDF multiplied by the number of options granted to CEO in Execucomp. Similarly, I estimate the number of non-scheduled options granted by multiplying the total options in Execucomp by the percentage of options that I estimate are non-scheduled from the IFDF. Finally, I use equation (3.1) to test whatever price targets affect scheduled or unscheduled options more. Table 3.5 summarizes the results.

The dependent variable is the number of options granted and not their value. In general, the number of unscheduled options responds positively to price targets while there is no significant association between price target and the number of scheduled options. The total number of options granted (TOT) is also positively associated with price targets in all specifications. The results also show that previous price targets are negatively associated with unscheduled options. As shown in chapter 2, price targets are positively related to future share prices. Thus, the forecasted increase of previous year share price boosted the value of unscheduled options making them worth more. Besides, even the number of unscheduled options. On the contrary, the number of scheduled options granted last year is positively associated with today's scheduled options. This might be due to the fact that scheduled options are a fixed component of CEO pay, and thus might be updated year after year as a sort of length of service bonus (as it is confirmed by the positive relation with CEO's age).

Table 3.5: Multivariate analysis of the effects of price targets on options from 1999 to 2010. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%, 5% and 1% levels respectively. The *t*-statistic are reported in parentheses.

•	(1)	(2)	(3)	(4)	(5)	(6)
	TOT	SCH	UNSCH	TOT	SCH	UNSCH
EXPROFIT	1284.5**	91.77	842.8**	-368.2	387.1	757.0**
	(2.48)	(0.34)	(2.03)	(-0.45)	(1.60)	(2.16)
EXPROFIT_1	-1096.9**	1.418	-819.9*	-374.7	-180.3	-951.7***
	(-2.25)	(0.01)	(-1.91)	(-0.48)	(-0.72)	(-2.62)
TOT_1	-0.805***					
	(-141.79)					
SCH_1		0.229***				
		(14.79)				
UNSCH_1			-0.0463***			
			(-2.94)			
RETURN	-41055.7***	-5037.3	-29899.1**			
	(-2.93)	(-0.51)	(-2.00)			
SALE	-84266.4***	-24541.3	-47249.7			
	(-3.11)	(-1.28)	(-1.62)			
CSHR	742.2***	124.9	-203.9*			
	(5.41)	(1.63)	(-1.75)			
EFFICIENCY	28096.7***	2441.1	18136.2***			
	(6.54)	(0.89)	(4.35)			
FEMALE	-169405.2***	33864.4	-254362***			
	(-2.83)	(0.79)	(-3.89)			
AVGAGE	300.1	888.1**	598.6			
	(0.56)	(2.37)	(1.05)			
AT	40370.1	32623.5*	-12981.4			
	(1.48)	(1.68)	(-0.44)			
YEAR	yes	yes	yes	yes	yes	yes
DUMMIES						
COMPANY	yes	yes	yes	yes	yes	yes
DUMMY						
CONSTANT	738103.9***	-18984.5	629615.4***	261507.0***	143845.8***	58932.6***
	(4.66)	(-0.16)	(3.55)	(7.96)	(14.27)	(4.04)
Ν	11439	5513	5513	12901	7692	7692

### **Appendix 3.B: Dot.com companies**

In the chapter I use companies that are reported in IBES, Execucomp and IFDF datasets. IFDF is required to obtain information on stock options, like the amount scheduled or

unscheduled options. However, Execucomp includes other firms that do not appear in IFDF (some of them because do not offer stock-based compensation).

In this appendix I study the substitution between cash-based compensation and stockbased compensation in the expanded dataset, which includes all firms in Execucomp for which IBES reports price targets.



As shown by the Venn diagram, in the chapter I focus on firms that result from crossing all three dataset: IBES, IFDF and Execucomp. Execucomp covers 3,168 companies, and roughly 87% (2,742) of them are also covered by IBES. Among such companies, IFDF covers approximately 71% (1,946) of them.

As in text regressions, the expected profit is the yearly company average of the difference between the price announced and the share price at the announcement days.

The results reported in table 3.6 show that by including all companies in Execucomp for which IBES provide price target the substitution effect between cash-based compensation and stock-based compensation disappears.

	Execucomp Companies					
	(1)	(2)	(3)	(4)		
	SBC	SBC	SBC	SBC		
EXPROFIT	0.0000114	0.0000091	0.00000906	-0.00000725		
	-0.53	-0.42	-0.42	(-0.33)		
SBC_1			0.0279***			
			-3.08			
EXPROFIT_1				0.0000688***		
				-3.18		
RETURN	-0.0373***	-0.0370***	-0.0370***	-0.0343***		
	(-7.39)	(-7.35)	(-7.33)	(-6.33)		
SALE	0.00277	-0.0319***	0.00235	0.000932		
	-0.38	(-3.08)	-0.32	-0.12		
CSHR	0.0000276	0.0000122	0.0000316	0.0000389		
	-0.5	-0.22	-0.58	-0.7		
EFFICIENCY	0.00290**	0.00370***	0.00247*	0.000707		
	-2.15	-2.73	-1.83	-0.43		
FEMALE	0.0202	0.0246	0.0188	0.0181		
	-0.87	-1.06	-0.82	-0.76		
AVGAGE	0.000811***	0.000814***	0.000803***	0.000747***		
	-4.06	-4.08	-4.01	-3.62		
AT		0.0496***				
		-4.75				
YEAR DUMMIES	yes	yes	yes	yes		
COMPANY DUMMY	yes	yes	yes	yes		
CONSTANT	0.404***	0.277***	0.396***	0.409***		
	-6.99	-4.67	-6.8	-7.27		
N	13666	13666	13570	13031		

Table 3.6: Multivariate analysis of the effects of price targets on Execucomp firms and on Execucomp firms excluded dot.com companies from 1999 to 2010. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%, 5% and 1% levels respectively. The t-statistic are reported in parentheses.

· · ·	/					
	Excluded dot.com companies					
	(5)	(6)	(7)	(8)		
	SBC	SBC	SBC	SBC		
EXPROFIT	0.000561***	0.000506***	0.000534***	0.000624***		
	-3.62	-3.26	-3.45	-3.13		
SBC_1			0.0252***			
			-2.67			
EXPROFIT_1				-0.0000969		
				(-0.51)		
RETURN	-0.0380***	-0.0378***	-0.0380***	-0.0397***		
	(-7.07)	(-7.03)	(-7.05)	(-7.01)		
SALE	-0.0113	-0.0429***	-0.0105	-0.0118		
	(-1.47)	(-3.99)	(-1.36)	(-1.49)		
CSHR	0.000018	0.00000461	0.0000219	0.0000295		
	-0.33	-0.09	-0.41	-0.54		
EFFICIENCY	0.000216	0.001	0.0000718	0.00104		
	-0.1	-0.47	-0.03	-0.47		
FEMALE	0.016	0.0185	0.0145	0.0145		
	-0.66	-0.77	-0.6	-0.59		
AVGAGE	0.00100***	0.00100***	0.000988***	0.000901***		
	-4.96	-4.96	-4.88	-4.31		
AT		0.0458***				
		-4.19				
YEAR DUMMIES	yes	yes	yes	yes		
COMPANY DUMMY	yes	yes	yes	yes		
CONSTANT	0.476***	0.366***	0.460***	0.486***		
	-8.37	-5.83	-8.03	-8.21		
N	12691	12691	12605	12103		

Table 3.6 (continued).

This result appears to be due to companies in the dot.com sector. Excluding those firms, the evidence shows that price targets are significantly and positively associated with stock-options compensation.<sup>33</sup> Companies in the dot.com sector differ from other companies by two main aspects. First of all, as Cowen *et al.* (2006) point out, dot.com companies received overconfident optimistic price targets in the past, which were proven to be wrong once the dot.com bubble collapsed. Thus, investors might have lost the confidence in analysts' price

<sup>&</sup>lt;sup>33</sup> I proceed as follow. I use the 4-digit SIC-codes to assign each firm to one of the 49 industries as defined by French and Fama. I drop companies working in the "Computer Software" sector and finally run the same regression as specified in equation (3.1).

targets for such companies.<sup>34</sup> Second, many dot.com firms paid their managers with stock compensation during the "start up" period, since they have liquid issue.<sup>35</sup> That is, because of the nature of the sector that requires comparable high human capital but has relative fewer entrance barriers, firms financed themselves by paying employees with stocks rather than (non-existent) cash. In this case, the trade-off between cash and stock cannot be adopted as strategy. Indeed, I do not find any significant association among price targets and the substitution effect for dot.com firms.

<sup>&</sup>lt;sup>34</sup> Investors might look suspicious at analysts' price target in software sector because of the dot.com collapse, despite I find that price targets are actually reliable (using the same procedure enlightened in chapter two). <sup>35</sup> Murphy (2003) uses Execucomp data between 1992 and 2001, finding that firms classified as "new economy"

used stock option compensation more intensively than large firms.

## Chapter Four: Stock Options: the End of Backdating

#### 1. Introduction

Chapter one showed that the use of stock options as a form of compensation increased in the 1990s to become the largest single component of CEOs compensation in 2000/2001. According to the principal agent model, the more options the CEO holds, the more she/he is incentivized to raise the share price above the strike price, which aligns the CEO's incentives with that of shareholders.

Stock options give the right to buy company stock at a pre-established strike price or exercise price. In US, stock options are usually granted at the money, so the strike price is the stock price at the grant date. As the Wall Street Journal (Forelle *et al.*, 2006) pointed out, granting options in the money (i.e. when the strike price is below the share price) is not illegal per se but granting options in the money is a cost in term of accounting rules that reduces the profit and thus must be included into the company books. Also, in most cases the options plans approved by shareholders explicitly requires that options be set at the money. A violation of such condition could result in an allegation of security fraud (Forelle *et al.*, 2006).

When options are granted at the money, executives can manipulate the timing of the grants in order to inflate options' value. Yermarck (1997) showed that company share price tended to drop before options granting and increase afterward. Aboody *et al.* (2000), Chauvin *et al.* (2001) argued that these abnormal patterns could be due to a manipulation of information by officers. For instance, a CEO might withhold good news or release bad news before she/he is awarded options - a practice known as spring loading. The subsequent fall in the share price will guarantee to the receiver a lower strike price. Subsequently, a CEO might release positive information right before exercising her/his options. Although these forms of timing

options/manipulating news announcements would be illegal as insider trading if the person was to purchase or sell shares, it is viewed as legal in the granting or selling of options, though many regard it as amoral.

Frydman *et al.* (2010) shows that spring loading alone cannot fully explain the abnormal return around option grants nor can "luck" (Lie, 2005). Lie (2005) shows that this abnormal pattern that most likely reason executives have historically made abnormal returns from options is that the grant date was opportunistically chosen retroactively - a process known as backdating. Backdating occurs when managers choose retroactively a past date as a grant date when the share price was particularly low in order to have a favorable strike price. Specifically, officers that backdate options go backward through the calendar looking for a date in which the share price was particularly low. Once found, they announce the options as granted at the money that day. As a result, they are able to obtain an option that looks granted at the money but de facto is awarded in the money. Lie (2005) and Heron *et al.* (2007) have used the abnormally low prices of shares at the dates when firms grant stock options to show that backdating was frequent in the 1990s.

In 2002 the U.S. Senate introduced the Sarbanes–Oxley Act (SOX) in an attempt of restore the trust in the market that the Enron and other corporate and accounting scandals had destroyed along with millions of dollars of shareholder wealth. Among others countermeasures, the SOX required that firms report to the SEC the options granted to insiders within two business days, which if followed would have eliminated backdating though not spring loading. Before August 2002, insiders had to file Form 4 and submit to the SEC "within ten days after the close of the calendar month in which the transaction had occurred" (Brochet, 2010, p. 420). In 2006, the SEC implemented the rules to increase the transparency of granting options. They required that top officers disclose if they are "timing options grants to make them more lucrative to executives and other employees" (Bickley, 2008, p. 15). Firms had to report to the SEC the price at the grant days, the grant date, if the strike price was lower than the share price

as well as to give reasons for having chosen a particular date as a grant date (Bickley, 2008).

In this chapter I examine the effect of the SOX and its further implementation in 2006 on the extent of backdating of options. I find that it effectively ended this form of giving options to executives. I then examine the extent to which firms substituted spring loading strategies by strategically timing their options granting around price targets announcements as an alternative way to assure that executives profited from options. The evidence shows that price targets have increasingly played a crucial role in the options granting process. Firms granted an increased number of options before analysts' gave positive price target announcements. The change in the mode of granting options is more consistent with the managerial power model than the arm's-length bargaining model of setting executive pay. Before 2002, an arms-length bargaining board might consider granting backdated options after a negative price targets announcement to avoid having underwater options that would reduce the motivation of CEOs and other executives. After 2002, an the ideal board might still grant options after a pessimistic price target caused by factors beyond the CEO's control. But a board dominated by management would behave differently. By definition a managerial power board always favors a strategy that enriches the CEO. Before 2002, it would backdate options if the share price were expected to rise while switching to other forms of compensation if the share price is expected to fall. In 2002, once the SOX took place, such a board would switch from backdating options to granting options before a positive price target announcement but not before a negative price target announcement. The empirical evidence shows that boards behaved according to the managerial strategy model.

## 2. Data and methodology

I obtained the official grant date, the exercise date, the expiry date and the exercise price of insider's grants from IFDF. That information allows me to estimate the percentage of options backdated as described in chapter one. I focus on transactions involving CEOs for the fiscal years 1999 to 2010.

I obtained price target data from IBES, where price target is defined as "the projected price level forecasted by an analyst within a specific time horizon" (Glushkov, 2009, p. 6). Price forecasts are available from March 1999. I retained only 12 months horizon forecasts, which constitute almost 98% of the entire dataset. Chapter 2 of this dissertation gives the details of the dataset and show that price targets have predictive value about future share prices and that the market responds to them in the short run, thus viewing them as valid information. I merge IFDF information with IBES Detail History Price Target file. For every option granted I determine whether IBES reports a price target announcement seven days before or after the grant - a 14-day window.<sup>36</sup> If more than one announcement was done within this window, I retain the announcement closest to the grant date. I report the price target announced and the announcement day. I create a dummy (BEFORE) taking the value of one if a stock option was granted from one to seven days before a price target announcement and zero if it was granted after or no price targets were announced in the 14-days window. Similarly, I create a true/false variable (AFTER) taking the value of one if a stock option was granted zero to seven days after an announcement, and zero if no announcement was reported within 14-days window or a price target was announced before. Finally, I create a dummy (AROUND) taking the value of one if the option was granted seven days after or seven days later or at the announcement date and zero otherwise.

<sup>&</sup>lt;sup>36</sup> In the appendix I provide evidence for a window of 28 days.

Figure 4.1 shows the distance between the announcement date and the grant date (in days). Fifty per cent of options granted around price targets are awarded before or at the price target announcement day. Sixty percent of options granted at the announcement day are unscheduled.



Figure 4.1: Difference between grant date and price target announcement day.

Table 4.1 reports the percentage of options granted before or at the announcement day, after the announcement day or independently by which I mean outside the window covering the price target announcement. The number of options granted independently of a price target announcement decreased significantly in the 2000s through 2009. In 1999 about 90% of options were granted independently from analysts' announcements, while in 2009, it dropped to 46%. Though the percentage of options granted independently increased in 2010, the 59% was still far below the 1999 level. Columns 2 and 3 show that the practice of granting options after price targets announcement increased particularly sharply. The remaining columns show that the increase is dominated by grants being given after announcements that foresee a rose in share price. The options granted before a positive price target announcements doubled from 36% to 63% in ten years, while the options granted after a positive announcements fell from

64% to less than 37%. Similarly, the options granted before a negative announcement decreased from 65% to 47% from 1999 to 2007.

V	Tetel Outlens			Increase in	share price	Decrease in	share price
rear	Total Options			expected (79%	% of grants)*	expected (219	% of grants) *
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	before	after	indip.	before	after	before	after
1999	5.65%	5.16%	89.19%	35.69%	64.31%	64.88%	35.12%
2000	12.55%	15.80%	71.64%	53.36%	46.64%	58.87%	41.13%
2001	11.24%	15.79%	72.97%	55.41%	44.59%	71.70%	28.30%
2002	10.53%	19.08%	70.39%	65.22%	34.78%	62.73%	37.27%
2003	13.03%	17.82%	69.15%	57.64%	42.36%	58.33%	41.67%
2004	15.43%	15.63%	68.94%	48.92%	51.08%	52.97%	47.03%
2005	13.31%	17.39%	69.30%	58.06%	41.94%	52.91%	47.09%
2006	14.11%	24.16%	61.73%	65.21%	34.79%	52.38%	47.62%
2007	14.84%	24.23%	60.93%	67.88%	32.12%	46.53%	53.47%
2008	14.38%	25.51%	60.11%	63.28%	36.72%	73.44%	26.56%
2009	10.81%	43.27%	45.92%	81.31%	18.69%	72.78%	27.22%
2010	15.12%	25.74%	59.15%	63.11%	36.89%	70.97%	29.03%

Table 4.1: Stock options granted before/after or independently from price targets announcements.

\*In 0.48% of cases options are granted around an announcement date for which I do not have information on share price at the announcement days.

To isolate the effect of the change in the law on the relation between price target announcements and granting options, I generated a dummy variable called "SOX" that assumes the value of 0 if the transaction occurred before 2002 and 1 if it was granted during 2002 or later. This variable is designed to catch the effect of change in the reporting rules as prescribed by the new law taking place in 2002. Likewise, I also create a variable defined "SEC" that it takes the value of the unity if the options were granted after 2006, when the SEC strengthened the SOX criteria.

## 3. Backdating

This section tests the effect of the change in reporting rules in preventing backdating in the options granting process, and if this had an unanticipated effect in reinforcing spring loading. The evidence tells a clear story. Introduction of the SOX in 2002 reduced backdating. Before 2002 boards frequently backdated options by strategically picking a grant date in the past on which the share price was particularly low. This backward strategy allowed CEOs reporting a favorable low strike price, and by so, inflating the option's value. After 2002, when the SEC required that insiders report the options they were granted within two days, backdating disappeared. But firms found another way to "inflate" the value of executives' options by using analysts price targets, a forward-looking strategy.

#### **3.1 Multivariate analysis**

To estimate the effect of the change in the law on firm option-granting behavior, I run the OLS regression with fixed effects on the following specification:

$$DEP_VAR_{i,t} = \propto +\beta_1 SOX_{t,i} + \beta_2 SEC_{t,i} + \beta_3 X_{t,i} + \delta yr * +\lambda_i + \varepsilon_{i,t}$$
(4.1)

where  $DEP_VAR_{i,t}$  stands for dependent variables. The dependent variables are set as follow:

• BD: is a dummy taking the value of one if options granted were backdated and zeros otherwise. An option is defined as backdated if the share price at the grant date is the lowest in a window of 40 days. As described in chapter one, I define the transaction share price (and day) as the one that minimizes in absolute value the distance between the strike price and the stock closing price at the reported transaction date or within two

days before the transaction. This because the SOX required in 2002 that the transaction be reported to the SEC within two business days.<sup>37</sup>

• PR: is a continuous variable measuring the probability that an option was backdated or spring-loaded. To do it, I construct a symmetric window of 20 trading days before and after the grant. For each of these 40 days I report the share closing price. I then count how many times the share price at the grant date is actually equal or lower than any share price reported in the 40 days window.<sup>38</sup> For instance, according to the previous definition, backdated options are those set at the minimum of a share price valley. In this case the number of time the stock closing price at the grant date is equal or lower than the stock closing price in the window of 40 days (20 days before and 20 days later) is 40/40, so 100%. On average I expect that the closing price at the grant date is 50% higher and 50% lower than the share prices of the transactions occurred in the window of 40 days. This variable is meant to catch all those transactions that are not necessarily at the minimum of a share price valley but still were granted at favorable lower share price. For instance, a board might decide to set the options at the second lowest share price in order to look less suspicious.

It is more difficult to use spring loading to grant stock options at the bottom of the V share price valley. A share price that is falling might suddenly increase before falling again (whisk of the tail), so that the CEO/board that intends to spring load might

<sup>&</sup>lt;sup>37</sup>Since in one single day more than one transaction may occur and each transaction may report a different strike price, it might be possible that a grant is assigned to a day (with its relative share price) while another transaction is assigned to a different day. Therefore, it is possible that one or more transactions reported by a company are considered backdated, while others reported by the same company are not considered backdated even if these transactions occur in the same day. For those 368 transactions occurred in 45 companies in 46 days, I retain the daily average of the true/false variable defining a transaction as backdated. The same issue also emerges on the variable PR and therefore I adopt the same methodology by retaining only the daily average.

<sup>&</sup>lt;sup>38</sup>The procedure I adopt is similar but different to the one described by Edelson et al. (2009). Instead of looking whether the stock closing price at the grant is lower or higher than the stock closing price of the symmetric window of trading days around the grant, the authors ranked the stock closing price of a symmetric window around the grant day and test whether the stock price at the grant "has equal probability of receiving any rank within the window" (Edelson et al., 2009, p. 3).

grant the options too early or too late and not catch the exact minimum. Nonetheless, despite the options are not at the lowest share price, the return is still favorable.

- BEFORE: is a dummy taking the value of one if a stock option was granted when a price target announcement was made or within seven days before, zero otherwise.
- AFTER: is a dummy taking the value of one if a stock option was granted from one to seven days after a price target announcement and zero otherwise.
- AROUND: is a dummy taking the value of one if the option was granted seven days after or seven days later or at the announcement date, zero otherwise.

The independent variables are:

- SOX: is a dummy that assumes the value of 0 if the transaction occurred before 2002 and it takes the value of 1 if it was granted during 2002 or later.
- SEC is a dummy that takes the value of the unity if the options were granted after 2006, when the SEC strengthened the SOX criteria, and zero otherwise.

 $X_{i,t}$  is a vector of observed individual characteristics:

- SALE: is the natural logarithm of the gross sales or of the amount of billing for regular sales in thousand of dollars.
- AT: is the natural logarithm of the current assets in million.
- CHSR: is the number of shareholders of ordinary/common capital (in thousands dollars).
- EFFIC: is the difference between Tobins Q at time *t* and Tobins Q at time (*t-1*), and it is meant to catch the need of restoring efficiency.

- FEMALE: a dummy variable taking the value of one if the CEO is a woman. If more than two CEOs are reported in a firms it stands for the percentage of female as CEO.
- AVGAGE: indicate the age of the CEO. If more than one CEO is reported in a company, the item stands for the average age of the officers.
- $yr^*$  stands for time dummies for each fiscal year from 1999 to 2010.
- $\lambda_i$  is firm fixed effect.

I summarize the results of the OLS regression with fixed effects in table 4.2.

Table 4.2	: Multivaria	ate analysis of	the effects o	f law on ti	iming strategy	y of	granting o	options	from	ı 1999
to 2010.	Estimates	are reported.	• *, **, ***	<sup>i</sup> indicate	significance	at	10%, 5%	and	1%	levels
respective	ely. The t-st	atistic are rep	orted in par	entheses.						
	14	x		(2)		1		( = )		

	(1)	(2)	(3)	(4)	(5)
	BD	PR	BEFORE	AFTER	AROUND
SOX	-0.0374***	-0.0599***	0.100***	-0.00299	0.0971***
	(-4.46)	(-4.62)	(4.82)	(-0.16)	(4.43)
SEC	-0.00193	0.0332**	0.0830***	0.0235	0.106***
	(-0.21)	(2.36)	(3.67)	(1.18)	(4.47)
RETURN	0.00818**	0.00124	-0.0249**	-0.00933	-0.0343***
	(2.01)	(0.20)	(-2.47)	(-1.05)	(-3.22)
SALE	-0.000890	-0.00615	0.0266	0.0180	0.0446**
	(-0.12)	(-0.55)	(1.48)	(1.13)	(2.35)
AT	-0.00956	-0.00353	0.0325*	0.0380**	0.0705***
	(-1.30)	(-0.31)	(1.78)	(2.36)	(3.66)
CSHR	-0.00000519	0.000000450	2.97e-08	-0.0000179**	-0.0000179**
	(-1.58)	(0.09)	(0.00)	(-2.50)	(-2.08)
EFFICIENCY	-0.00301***	-0.00330**	-0.0000250	0.00287	0.00284
	(-2.91)	(-2.06)	(-0.01)	(1.27)	(1.05)
FEMALE	0.00304	0.0497*	0.0242	-0.00708	0.0171
	(0.16)	(1.68)	(0.51)	(-0.17)	(0.34)
AVGAGE	-0.0000240	-0.00000283	-0.000463	-0.000179	-0.000642
	(-0.15)	(-0.01)	(-1.20)	(-0.53)	(-1.58)
YEAR	yes	yes	yes	yes	yes
DUMMIES					
COMPANY	yes	yes	yes	yes	yes
DUMMY					
CONSTANT	0.135***	0.649***	-0.232**	-0.244**	-0.477***
	(3.04)	(9.42)	(-2.11)	(-2.51)	(-4.10)
Ν	10748	10748	10781	10781	10781

Table 4.2 gives evidence that backdated options (BD) responded negatively to the introduction of SOX in 2002 and the implementation of the rule in 2006. The model also shows that the probability of backdating/spring loading (PR) fell with the introduction of SOX in 2002 but then increased with implementation of the rules. One possible explanation for this pattern is the following. Before 2002 managerial-dominated boards backdated options in order to inflate

the options' value. The introduction of the law in 2002 effectively reduced the phenomenon but set CEOs/the boards looking for alternative strategies, such as spring loading. In 2006, the introduction of a new law did nothing to discourage the already tested practice of spring loading, so more boards chose this way to enrich their executives. Supporting the interpretation that the changes reflected the operation of a management-dominated board the changes in 2002 and in 2006 augmented the options granted before a price targets announcements, but did not affect the practice of granting options after an announcement.

The SOX law thus produced the desired effect of eliminating the backdating phenomenon but reinforced spring loading strategy. Boards and CEOs shifted from backdating to the riskier but still lucrative forward-looking spring loading.

#### 4. Backdating, price targets and the law

Backdating is a strategy mainly associated with a managerial power model (which assumes that the board maximizes CEO pay under social constraints) rather than by an ideal board (arm's length bargaining model, which assumes the board of directors contract top executives pay in order to maximizes shareholders' interests). Still, it is possible that an armslength bargaining board could adopt backdating as a suitable strategy, as this reduces the riskiness associated with options and thus the risk premium that executives would want for shifting pay from cash to options.

### 4.1 Managerial dominated board vs. arms-length bargaining board

A managerial dominated board would always consider backdating as a way to increase the likelihood that options will profit management. But such a board might decide that when the share price is expected to fall, even backdated options are unlikely to pay off for executives. A backdated option is in the money at the transaction day but falls in the share price will reduce its intrinsic value. As argued in chapter three, such a board would likely switch from stock-based compensation to cash-based compensation. An arm-length bargaining board might consider granting backdated options if analysts expect a fall in the share price because of causes beyond CEO control. Granting options in the money when a share price is expected to fall would keep managers motivated ex post, once the share price actually fell. It is a way of indexing the strike price to the expected share price and thus providing an incentive to the CEO to beat that expectation. If the share price is expected to fall by 10% and the board has backdated the option to a period when share prices were 10% lower, it has effectively contracted for an appropriate at the money option. If the CEO manages to take actions that reduce the fall in the price to 5%, he or she will be rewarded by the share price exceeding the strike price on the backdated award.

Thus, the likely course of the share price determines whether which sort of board will backdate options. When share prices are expected to increase, the management-dominated board will backdate while the arms-length bargaining board will not. When share prices are expected to fall, the arms-length board will use backdating to maintain incentives while the management-dominated board is more likely to shift compensation to cash.

In 2002, the introduction of the SOX essentially eliminated backdating but allowed boards to shift toward a more forward-looking strategy, such as spring loading, for options. Spring loading can also be a desirable strategy in both models of board behavior depending on the sign of options' intrinsic value. A board might take advantage of the fact that the market reacts positively to analyst's price target announcements. An option granted at the money before a positive announcement will be most likely in the money soon after, whereas an option granted before a negative announcement will most likely be out of the money soon afterwards. A management dominated board would consider granting an option before a positive price target announcement so that the option is in the money after the announcement. It makes the managers wealthier. Granting options after a negative announcement would produce a lower strike price soon after the announcement but would risk the share price falling more in the future, depending on how the market responded to the price target announcement.

An arms-length bargaining board would also look at the price target to infer the likely value of options and try to set the right options compensation according. An ideal board would most likely adopt as a dominant strategy granting options after an announcement, so that the share price (and the strike price) fully reflect the new information and the CEO would neither be paid for "luck" (in case of a positive announcement) nor for "misfortune" (in case of a negative announcement) if the stock price curse is beyond his/her control. Table 4.3 summarizes the strategies considered by both models.

To test which models is predominant I proceed as follow. For each company and year, I compute the total number of backdated options granted before a positive price targets announcement ( $Backdated_{B,P}$ ) as well as the total number of backdated options granted after a negative announcement ( $Backdated_{A,N}$ ). I then take the difference between backdated stock options granted before or at the positive announcement date and backdated stock options granted after a negative price target announcement. Under the null hypothesis of an ideal armslength bargaining board the difference should be negative. The management dominated board would by contrast produce a positive difference.

From the previous multivariate analysis we know that the introduction of SOX reduced the backdating phenomenon but increased the spring loading strategy. For testing which model better explains the spring loading phenomenon, I compute for each company and year the total number of options granted before or at a positive price target announcement ( $Before_p$ ) and the total number of options granted after a positive announcement ( $After_p$ ). I repeat the procedure for options granted before ( $Before_N$ ) or after ( $After_N$ ) a negative announcement. I test whether the difference between the numbers of options granted before and after a positive (negative) price announcement is greater or less than zero. The hypothesis of an ideal board predicts that the number of options granted before a positive price target announcement is less than the number of options granted after a positive price target announcement. Also, in the case when analysts expect a drop in the share price, the ideal board would grant options after the announcement while the management dominated board will consider granting options after the announcement or granting no options and shifting pay into cash compensation. Thus the ideal board and the rent-seeking board may choose the same strategy when the share price is expected to fall. But the ideal arms-length board will grant fewer options after a negative price target announcement than after a positive price target announcement than a management dominated board.

	Ideal-board of directors		Rent-seeking board		
Expected intrinsic value	Positive	Negative	Positive	Negative	
Granting strategy before 2002 or 2006	No backdated options, grant options after announcements.	Backdated options. If the causes beyond the expected drop in share are out of CEO's control, the board can grant backdated options so once the price drops options are ideally at the money and manager still incentivized. Grant after announcements.	Backdated options and grant options before announcements.	Grant options after or switch to other remuneration forms, e.g. cash- based compensation since the options will be at the money or out of the money once the forecasted dropping trend ends.	
Granting strategy after 2002 or 2006	No backdating. Grant after the price target announcement, so the strike price reflects the new information and managers are not rewarded for "luck".	No backdating. Grant extra options after the price target announcement, so the strike price reflects the new information and managers are not punished for "misfortune" and also incentivized when most needed.	Grant options before the announcements, so they are in the money soon after. No backdating.	Grant options after or switch to other remuneration forms, e.g. cash- based compensation, since the options will be at the money or out of the money once the forecasted dropping trend ends.	

Table 4.3: Different strategy adopted by both ideal and rent-seeking board before and after the introduction of the SOX according to different price targets scenarios (positive vs. negative).

I summarize in table 4.4 the t-tests.

The empirical evidence rejects for all three scenarios the hypothesis of an ideal board in favor of the alternative hypothesis of a rent-seeking board. It rejects the hypothesis that CEOs receive more stock options after a positive announcement (as predicted by an ideal board model) in favor of the alternative (as predicted by a rent-seeking model) and rejects the hypothesis that CEOs are awarded more stock options after a negative announcement.

#### Table 4.4: t-tests.

	$H_0$ : Ideal	$H_A$ : Rent-	Obs.	Mean	Std. Dev.	t	$H_A$ : mean
	board	seeking board					> 0
		alternative					
$Backdated_{B,P}$	mean ≤0	mean >0	25510	894	33364	4.28	Pr(T > t) =
$-Backdated_{A,N}$							0.0000
$Before_P - After_P$	mean ≤0	mean >0	25510	17623	812300	3.47	Pr(T > t) =
							0.0003
$Before_N - After_N$	mean ≤0	mean >0	25510	3362	173512	3.09	Pr(T > t) =
							0.0010

## 4. Conclusion

This chapter explored how the introduction of the Sarbanes-Oxley Act in 2002 and its further implementation in 2006 affected the supply of options. I show that the change in the law succeeded in reducing the practice of backdating options but it led some board to substitute the practice of backdating options with a more forward-looking strategy of granting options around analysts' price targets announcements (spring loading). It also showed show that price targets have become increasingly linked to the granting options process. The number of options granted independently from analysts' price targets announcements declined sharply from 1996 to 2010. Among the options granted around price targets announcements 79% were awarded when analysts expected an increase in the share price and the vast bulk were granted before the positive announcement. Finally, I discuss the options granting timing strategies that could be

adopted by both an ideal board and a rent-seeking board before and after the introduction of the SOX. I argue that what makes backdated options/spring loaded options a rent-seeking strategy rather than an ideal strategy is the timing of the grant and the sign of the future expected share price. I give evidence that the results reject the hypothesis of an ideal board in favor of a rent-seeking board alternative and I argue that a rent-seeking model rather than an arm-length bargaining model better explains the emerging evidence.

## Appendix 4.A

This appendix examines the robustness of my results. Specifically, in the body of the chapter I considered an option granted before (after) a price targets announcements awarded within seven days before (after) a price target announcements. In this appendix I show that the results hold if I expand the window from fourteen to twenty-eight days (fourteen days before, fourteen days after). Table 4.6 gives the regressions for the expanded window.

Table 4.6: Multivariate analysis of the effects of law on timing strategy of granting options from 1999 to 2010. Estimates are reported. \*, \*\*, \*\*\* indicate significance at 10%, 5% and 1% levels respectively. The *t*-statistic are reported in parentheses.

	Entire dataset					
	(1)	(2)	(3)			
	BEFORE	AFTER	AROUND			
SOX	0.140***	0.00430	0.144***			
	(6.34)	(0.21)	(7.01)			
SEC	0.0851***	0.0165	0.102***			
	(3.55)	(0.75)	(4.54)			
RETURN	-0.0285***	-0.0000731	-0.0286***			
	(-2.66)	(-0.01)	(-2.86)			
SALE	0.0102	0.0191	0.0293			
	(0.53)	(1.08)	(1.64)			
AT	0.0403**	0.0254	0.0657***			
	(2.08)	(1.42)	(3.63)			
CSHR	0.0000012	-0.00003***	-0.00003***			
	(0.14)	(-3.33)	(-3.15)			
EFFICIENCY	-0.000220	0.00208	0.00186			
	(-0.08)	(0.83)	(0.73)			
FEMALE	-0.0317	-0.00923	-0.0409			
	(-0.63)	(-0.20)	(-0.87)			
AVGAGE	-0.000334	-0.000228	-0.000563			
	(-0.82)	(-0.61)	(-1.48)			
YEAR	yes	yes	yes			
DUMMIES						
COMPANY	yes	yes	yes			
DUMMY						
CONSTANT	-0.112	-0.0974	-0.209*			
	(-0.96)	(-0.90)	(-1.92)			
N	10781	10781	10781			

## Conclusion

Research on executive compensation is divided between two views or theories of how boards determine the pay of CEOs and other top managers. One view is that CEO pay is the result of a contracting process between CEO aiming to maximizing her/his compensation and a board of directors aiming to maximize shareholders' interests (arm-length bargaining model, using the definition of Bechuck *et al.* 2005). The other view is that management dominates boards of directors so that the board seeks to maximize managers' compensation within constraints imposed by social costs and market penalties (managerial power model, using the definition of Bechuck *et al.* 2005). In the first case, the pay is the outcome of the intersection of market forces. It is optimal in the sense that the extra dollar gained by a CEO is equal to the marginal contribution of her/his work. In the second case, managers use private information to structure compensation for their benefit at the expense of shareholders and influence the board of directors through weaknesses in corporate governance to set their own pay and gain rents at the expense of shareholders. This line of thinking questions the efficacy of stock-based incentives to solve the principal agent models and views incentive schemes as part of the agency problem itself. The rise of inequality is no longer justified by a rise in efficiency.

Bebchuk *et al.* (2003) argue that the observed pay structure is most likely a compromise between market forces, which mitigates managerial rent-seeking, and managerial-power that favors top executive compensation. Reviewing empirical studies, Frydman *et al.* (2010) conclude "both managerial-power and competitive market forces are important determinants of CEO pay" but leave open the question of their relative importance. The first two chapters of this dissertation add new data for assessing the two hypotheses regarding determination of CEO pay in the U.S. The analysis in the next two chapters shows that the managerial power model fits better with the new evidence than the arm's length bargaining model. In chapter one, I study the changes in US CEO compensation from 1996 to 2010 using data from the Execucomp dataset and the IFDF dataset. I describe the methodology by which I matched Execucomp data from company's annual proxy statement (DEF 14A SEC) with data from the Insider Filing Data Feed on U.S. insider activities. I found that the number of options reported in the two datasets is highly correlated (except in 2006 when Execucomp changed the mode of reporting options) to allow to analyze both dataset combined. I show that CEO pay rose by \$2.7 millions between 1996 and 2010. I also show that the changes varied greatly over time. I find that most of the increase in stock-based compensation from 1996 to 2001 was due to increased grants of stock options, but that options became less popular toward the end of the period in favor of direct grants of stocks, which in 2010 constituted almost 28% of total compensation. Regarding the composition of options over time, I give evidence that the percentage of scheduled options increased by almost 25% between 1997 and 2010 and that backdated options disappeared after 2002. Cash-based compensation increased more modestly over the period.

The second chapter analyzes Institutional Broker Estimate System (IBES) Detail History Price Target data file, a dataset that records analysts' price targets for the share prices of companies, an indicator of the likely future change in a firms' stock price. Focusing on US firms I show that 98% of price targets have one-year horizon. In 18.65% of the cases a firm receives more than one price target per day. I also find that the number of price targets announcements issued by analysts is positively associated with company share price's volatility and I give evidence that price target is positively associated with forecasted price, and thus when analysts expect a rise (fall) in next year share price it is more likely to occur than not. Lastly, I provide evidence that positive (negative) announcements are followed by an increase (decrease) in share price, which implies that investors incorporate analysts' information.

The third chapter analyzes the impact of price targets, as reported by the Institutional Broker Estimate System in its Detail Price History Target database, on executive compensation. The share price that analysts expect for a firm in the future has not, to my knowledge been previously analyzed in studies of executive compensation. I find that analysts' price targets alter the composition of executive pay between cash-based compensation and stock-based compensation. When analysts forecast a rise in the share price for a firm, its compensation package tilts toward stock-based compensation. When analysts forecast a fall in the share price, the compensation package tilts toward cash-based compensation. This pattern is more readily explicable by the managerial power model than the arm's length bargaining model. Consistent with this interpretation, the trade-off is stronger in companies that have weaker corporate governance.

The fourth chapter explores the impact of the Sarbanes-Oxley Act and its implementation by the SEC affected the granting of stock options. I find that the change in the law induced CEOs to substitute the backward-looking practice of backdating with a forward-looking strategy of spring loading options using price targets. The chapter documents that the timing of granting options is increasingly linked to price target announcements, with the number of options granted in periods outside a window around analysts' price targets announcements declining sharply from 1996 to 2010. I also discuss the options granting timing strategies that an ideal board and a rent-seeking board would likely follow before and after the introduction of the SOX. The timing of granting options and the relation of the grants to analysts price target announcements load fits better with the rent-seeking managerial power model than the arms-length deal strategy.

To my best knowledge price targets have not been used in studies of executive compensation. The dissertation shows that they provide insights into the determinants of executive pay in favor of the rent-seeking model. Clearly, there remains large scope for further investigation on the relation of CEO pay and price targets. One area that deserves further analysis is the relation between scheduling options around regular price targets announcements or awarding unscheduled options after surprising price target announcements.

## Terminology

In this section I provide a brief terminology of the technical terms I used.

## A. Stock options glossary

Options are financial contracts that give the right to buy (call option) or sell (put option) an underlying asset by a specific date (expiration or maturity date) for a fixed price (exercise or strike price). The right does not constitute an obligation. The employee stock options (ESO) are call options on the stock of the company. Typically, the strike price of employees' call options is set equal to the firm' share price at the grant date and the employees have to wait a predefined period of time before they can exercise those options (vesting period). Usually, the number of calls options granted in a single transaction coincides with the number of shares underlined, but an option can underline more or less shares.

American	if they are exercisable at any time.	
At the money	if the exercise price is equal to the share price of the underlying stock.	
Backdated	if the grant date is manipulated so that the company dates the options granted to employees at earlier period when the share price was lower than the share price at the time the options were actually granted. In U.S. most options are granted at the money, thus backdating allows to report a strike price lower than the company's share price at the grant date. This procedure makes the options de facto in the money and thus inflates the value of the options granted to executives.	
Bermudian	if they can be exercised after a predefined period of time has passed (vesting period). These options are hybrid between European Options and American Options (Rubinstein, 1995).	
Europeans	if the exercise date coincides with the expiration date.	
In the money	if the strike price is lower than the share price of the underlying stock.	
Out of the money	if the strike price is higher than the share price of the underlying stock.	
Scheduled	if they are granted at regular time year by year.	
Spring-loaded	if options are granted before a positive news is released. This procedure inflates the value of the options since the market reacts to the news pushing the share price up and thus the options' value as well.	
Unscheduled	if they are granted at irregular time year by year.	

Call Options on stock are said to be:

# **B.** Execucomp terminology

Execucomp collects top executive compensation detailed information directly from company's annual proxy statement (DEF 14A SEC). Thus, the change in the SEC's compensation disclosure rules adopted in 2006 resulted in a discontinuity in the Execucomp data. Below, I summarize the main implications.

	Before 2006	After 2006
Bonus	Short-term incentives pay in cash.	Defines cash earned by officers who met criteria that were not disclosed.
Long-Term Incentive	LTIP are earned if CEO satisfies criteria measured over a time period longer than one year, typically three to six years (Larcker, 1983). Execucomp reports LTIP until 2006 when the SEC introduced a new terminology to isolate executive incentive compensation paid in cash and no longer require the distinction between short-term incentive and long-term incentive.	No longer reported.
Non-equity incentives		Terminology introduced by SEC in 2006 to define cash incentive awarded if executives meet a disclosed pre-established target. The new terminology shifted part of cash compensation previously reported under the label "bonus" and LTIP to the new label "non- equity incentive".
Stock options value	Before 2006, Execucomp provides its own fair award value estimates of the monetary worth of stock options granted using a non- standard Black-Scholes formula.	After 2006 SEC imposes companies to report the estimated fair value of the stock options granted in the proxy statement, and Execucomp decided to drop its method in favor of the company reports. The SEC allows companies to use different methods of evaluating options, including the Black-Scholes and the binomial options pricing model.
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